

Math 302.102 Fall 2010

Some Examples of a One-Dimensional Change of Variables

Suppose that  $X$  is a continuous random variable and that  $Y = g(X)$  for some continuous function  $g : \mathbb{R} \rightarrow \mathbb{R}$  so that  $Y$  is itself a continuous random variable. It is often the case in practice that one knows the density function of  $X$  and seeks the density function of  $Y$ . Fortunately, if  $g$  is a nice function (as it usually is in practice), then it is straightforward to determine the density of  $Y$  from first principles. Basically, one starts with the definition of the distribution function of  $Y$  substitutes in  $Y = g(X)$ , and solves for  $X$ . This produces an integral expression involving the density function of  $X$  which can then be differentiated using the fundamental theorem of calculus to yield the density function for  $Y$ . Sometimes this is called a *one-dimensional change of variables*. The following examples illustrate this technique. Remember that in order to use the fundamental theorem of calculus, it must be the case that a variable appears in the upper limit of integration and that no variable appears in the lower limit of integration.

**Example.** Suppose that  $X \sim \mathcal{N}(0, 1)$ . Let  $Y = e^X$ . Determine the density function of  $Y$ .

**Solution.** Let  $Y = e^X$ . For  $y > 0$ , the distribution function of  $Y$  is

$$F_Y(y) = \mathbf{P}\{Y \leq y\} = \mathbf{P}\{e^X \leq y\} = \mathbf{P}\{X \leq \log y\} = \int_{-\infty}^{\log y} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$

so that the density function of  $Y$  is

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} \int_{-\infty}^{\log y} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = \frac{1}{\sqrt{2\pi}} e^{-(\log y)^2/2} \cdot \frac{d}{dy} \log y = \frac{1}{y\sqrt{2\pi}} e^{-(\log y)^2/2}$$

for  $y > 0$ . *The random variable  $Y$  is an example of a log-normal random variable which is regularly encountered in the mathematical theory of stock option pricing.*

**Example.** Suppose that  $X \sim \mathcal{N}(0, 1)$ . Let  $Y = X^2$ . Determine the density function of  $Y$ .

**Solution.** Let  $Y = X^2$  so that

$$F_Y(y) = \mathbf{P}\{Y \leq y\} = \mathbf{P}\{X^2 \leq y\}.$$

Note that since  $X$  can take on *any* real value, we have

$$\begin{aligned} \mathbf{P}\{X^2 \leq y\} &= \mathbf{P}\{-\sqrt{y} \leq X \leq \sqrt{y}\} = \int_{-\sqrt{y}}^{\sqrt{y}} f_X(x) dx \\ &= \int_{-\sqrt{y}}^0 f_X(x) dx + \int_0^{\sqrt{y}} f_X(x) dx \\ &= \int_0^{\sqrt{y}} f_X(x) dx - \int_0^{-\sqrt{y}} f_X(x) dx \\ &= \int_0^{\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx - \int_0^{-\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx, \end{aligned}$$

and so

$$\begin{aligned}
 f_Y(y) &= \frac{d}{dy} F_Y(y) = \frac{d}{dy} \int_0^{\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx - \frac{d}{dy} \int_0^{-\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \\
 &= \frac{1}{\sqrt{2\pi}} e^{-(\sqrt{y})^2/2} \cdot \frac{d}{dy} (\sqrt{y}) - \frac{1}{\sqrt{2\pi}} e^{-(-\sqrt{y})^2/2} \cdot \frac{d}{dy} (-\sqrt{y}) \\
 &= \frac{1}{\sqrt{2\pi}} e^{-y/2} \cdot \frac{1}{2\sqrt{y}} + \frac{1}{\sqrt{2\pi}} e^{-y/2} \cdot \frac{1}{2\sqrt{y}} \\
 &= \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{y}} e^{-y/2}
 \end{aligned}$$

for  $y > 0$ . Note that the random variable  $Y$  has a  $\text{Gamma}(1/2, 1/2)$  distribution, or equivalently,  $Y \sim \chi^2(1)$  and often appears in statistical inference.

**Example.** Suppose that  $X \in \Gamma(a, b)$  so that the density of  $X$  is

$$f_X(x) = \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx}$$

for  $x \geq 0$ . Let  $Y = 1/X$ . Determine the density function of  $Y$ .

**Solution.** Let  $Y = 1/X$ . For  $y > 0$ , the distribution function of  $Y$  is

$$\begin{aligned}
 F_Y(y) &= \mathbf{P}\{Y \leq y\} = \mathbf{P}\{1/X \leq y\} = \mathbf{P}\{X \geq 1/y\} = 1 - \mathbf{P}\{X < 1/y\} \\
 &= 1 - \int_{-\infty}^{1/y} \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx} dx
 \end{aligned}$$

so that the density function of  $Y$  is

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} \left( 1 - \int_{-\infty}^{1/y} \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx} dx \right) = \frac{b^a}{\Gamma(a)} y^{1-a} e^{-b/y} \cdot \frac{1}{y^2} = \frac{b^a}{\Gamma(a)} y^{-a-1} e^{-b/y}$$

for  $y > 0$ . The random variable  $Y$  is an example of an inverse gamma random variable with parameters  $a$  and  $b$  and is used primarily in Bayesian statistics though it sometimes finds applications in actuarial science.