

Statistics 252 Midterm #2 – March 3, 2006

1. (10 points)

- (a) Suppose that a random variable Y has density function $f_Y(y|\theta) = (\theta + 1)y^\theta$, $0 \leq y \leq 1$. Determine the Fisher information $I(\theta)$ for this random variable.
- (b) Let Y_1, \dots, Y_n be independent and identically distributed random variables each with the density $f_Y(y|\theta)$ given in (a). Determine $\hat{\theta}_{\text{MOM}}$, the method of moments estimator of θ .

- 2.** (7 points) Consider a random variable Y with density function $f_Y(y|\theta) = 2\theta^{-2}y$, $0 \leq y \leq \theta$, for some parameter $\theta > 0$. Use the pivotal method to verify that if $0 < \alpha < 1$, then

$$\left(\frac{Y}{\sqrt{1 - \alpha/2}}, \frac{Y}{\sqrt{\alpha/2}} \right)$$

is a confidence interval for θ with coverage probability $1 - \alpha$.

- 3.** (10 points) Recall that a discrete random variable Y is said to be Poisson with parameter $\theta > 0$ if the density (also called probability mass function) of Y is

$$f_Y(y|\theta) = \frac{\theta^y e^{-\theta}}{y!}, \quad y = 0, 1, 2, \dots$$

Recall further that if Y is Poisson(θ), then $E(Y) = \theta$ and $\text{Var}(Y) = \theta$.

- (a) It turns out that the same formula for the Fisher information can be used for discrete random variables. Show that if Y is a Poisson random variable with parameter θ , then $I(\theta) = \frac{1}{\theta}$.

For parts (b), (c), (d), and (e) below, suppose that Y_1, Y_2, \dots, Y_n are independent and identically distributed Poisson(θ) random variables.

- (b) Show that $\hat{\theta}_{\text{MOM}}$, the method of moments estimator of θ , is $\hat{\theta}_{\text{MOM}} = \bar{Y} = \frac{Y_1 + \dots + Y_n}{n}$.
- (c) Show that $\hat{\theta}_{\text{MOM}}$ is an unbiased estimator of θ .
- (d) Calculate $\text{Var}(\hat{\theta}_{\text{MOM}})$.
- (e) Explain why $\hat{\theta}_{\text{MOM}}$ must be the minimum variance unbiased estimator (MVUE) of θ .

4. (8 points)

- (a) Let Y_1, Y_2, \dots, Y_n be independent and identically distributed random variables each with mean p and variance $p(1 - p)$. Here p is an unknown parameter between 0 and 1. If

$$\hat{p} = \frac{Y_1 + \dots + Y_n}{n}$$

then it is known that \hat{p} is an unbiased estimator of p . Show that the maximum value of $\text{Var}(\hat{p})$, the variance of \hat{p} , occurs when $p = 1/2$.

- (b) Suppose that $\hat{\theta}_1$ and $\hat{\theta}_2$ are both unbiased estimators of θ . Explain how $\text{Eff}(\hat{\theta}_1, \hat{\theta}_2)$, the relative efficiency of $\hat{\theta}_1$ to $\hat{\theta}_2$, can be used to decide which of these two estimator is preferable.