

1. (a) Since

$$\int_1^{\infty} x^{-2} dx = -x^{-1} \Big|_1^{\infty} = 0 - (-1) = 1$$

we see that taking $c = 1$ makes f a legitimate probability density.

(b) Since

$$\int_1^{\infty} x^{-1} dx = \ln|x| \Big|_1^{\infty} = \infty - 0 = \infty$$

we see that there is no such c that makes f a legitimate probability density.

(c) Using integration-by-parts twice (see Prerequisite Review Handout) gives

$$\int x^2 e^{-x} dx = -x^2 e^{-x} - 2x e^{-x} - 2e^{-x}.$$

Therefore, since

$$\int_{-1}^1 x^2 e^{-x} dx = [-x^2 e^{-x} - 2x e^{-x} - 2e^{-x}] \Big|_{-1}^1 = [-e^{-1} - 2e^{-1} - 2e^{-1}] - [-e^1 + 2e^1 - 2e^1] = e - 5e^{-1},$$

we see that taking $c = (e - 5e^{-1})^{-1} = e/(e^2 - 5)$ makes f a legitimate probability density.

(d) As in (c), using integration-by-parts twice gives

$$\int_0^{\infty} x^2 e^{-x} dx = [-x^2 e^{-x} - 2x e^{-x} - 2e^{-x}] \Big|_0^{\infty} = 0 - (-2) = 2$$

and so we see that taking $c = 1/2$ makes f a legitimate probability density.

(e) Since

$$\int_{-\infty}^0 x e^x dx = [x e^x - e^x] \Big|_{-\infty}^0 = (0 - 1) - (0 - 0) = -1$$

by see that taking $c = -1$ makes f a legitimate probability density. Note that $x e^x \leq 0$ for $x \leq 0$. This means that if we multiply $x e^x$ by a negative number it will always be non-negative. Hence, $f(x) = -x e^x$ for $x \leq 0$ is a non-negative function that integrates to 1.

(f) Note that the function $x e^x$ assumes both positive and negative values when $-1 \leq x \leq 1$. This means that there is no single value of c that could be multiplied by $x e^x$ to make it strictly non-negative. Hence, there is no possible value of c that makes f a legitimate probability density.

2.

(a) We find

$$\mathbf{P}\{0 < X < 2\} = \int_0^2 f(x) dx = \int_0^2 7e^{-7x} dx = -e^{-7x} \Big|_0^2 = 1 - e^{-14}.$$

(b) We find

$$\mathbf{P}\{0 < X < 2\} = \int_0^2 f(x) dx = \int_0^2 xe^{-x} dx = [-xe^{-x} - e^{-x}] \Big|_0^2 = 1 - 3e^{-2}.$$

3. (a) Notice that we must necessarily have $0 < a < 4$. Since

$$\mathbf{P}\{X \leq a\} = \int_0^a \frac{1}{8}x dx = \frac{x^2}{16} \Big|_0^a = \frac{a^2}{16},$$

we find that in order for $\mathbf{P}\{X \leq a\} = 1/2$ we must have

$$\frac{a^2}{16} = \frac{1}{2}$$

implying that $a^2 = 8$. The restriction that $a > 0$ implies that the unique value of a such that $\mathbf{P}\{X \leq a\} = 1/2$ is $a = \sqrt{8}$.

(b) As in (a), we must necessarily have $0 < a < 4$. Since

$$\mathbf{P}\{X \geq a\} = \int_a^4 \frac{1}{8}x dx = \frac{x^2}{16} \Big|_a^4 = 1 - \frac{a^2}{16},$$

we find that in order for $\mathbf{P}\{X \leq a\} = 1/2$ we must have

$$1 - \frac{a^2}{16} = \frac{1}{4}$$

implying that $a^2 = 12$. The restriction that $a > 0$ implies that the unique value of a such that $\mathbf{P}\{X \geq a\} = 1/4$ is $a = \sqrt{12}$.

4. Recall that the distribution function F of a random variable is defined as

$$F(x) = \mathbf{P}\{X \leq x\}.$$

Since X is a continuous random variable, we know that $\mathbf{P}\{X = x\} = 0$ for any $x \in \mathbb{R}$ so that

$$\mathbf{P}\{X < x\} = \mathbf{P}\{X \leq x\} = F(x).$$

(a) We find

$$\mathbf{P}\{X \leq 1\} = F(1) = \frac{1}{8}(1)^3 = \frac{1}{8}.$$

(b) We find

$$\mathbf{P}\{0.5 \leq X \leq 1.5\} = \mathbf{P}\{X \leq 1.5\} - \mathbf{P}\{X < 0.5\} = F(1.5) - F(0.5) = \frac{1}{8} [(1.5)^3 - (0.5)^3] = \frac{13}{32}.$$

(c) Notice that we must necessarily have $0 < a < 2$. Since

$$\mathbf{P}\{X \leq a\} = F(a) = \frac{a^3}{8},$$

we find that in order for $\mathbf{P}\{X \leq a\} = 1/2$ we must have

$$\frac{a^3}{8} = \frac{1}{2}$$

implying that $a^3 = 4$. Thus, the unique value of a such that $\mathbf{P}\{X \leq a\} = 1/2$ is $a = 4^{1/3} = \sqrt[3]{4}$.

(d) As in (c), we must necessarily have $0 < a < 2$. Since

$$\mathbf{P}\{X \geq a\} = 1 - \mathbf{P}\{X < a\} = 1 - F(a) = 1 - \frac{a^3}{8},$$

we find that in order for $\mathbf{P}\{X \geq a\} = 1/4$ we must have

$$1 - \frac{a^3}{8} = \frac{1}{4}$$

implying that $a^3 = 6$. Thus, the unique value of a such that $\mathbf{P}\{X \geq a\} = 1/4$ is $a = 6^{1/3} = \sqrt[3]{6}$.

5. Let

$$\Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt$$

denote the distribution function of a normally distributed random variable. Hence, we find the following.

(a) $\mathbf{P}\{X > 1\} = 1 - \mathbf{P}\{X \leq 1\} = 1 - \Phi(1) = 1 - 0.8413 = 0.1587$.

(b) $\mathbf{P}\{X < 1\} = \mathbf{P}\{X \leq 1\} = \Phi(1) = 0.8413$.

(c) $\mathbf{P}\{X \leq 1\} = \Phi(1) = 0.8413$.

(d) $\mathbf{P}\{-1 \leq X \leq 1\} = \mathbf{P}\{X \leq 1\} - \mathbf{P}\{X < -1\} = \mathbf{P}\{X \leq 1\} - \mathbf{P}\{X \leq -1\} = \Phi(1) - \Phi(-1) = (0.8413) - (1 - 0.8413) = 0.8413 - 0.1587 = 0.6826$.

(e) $\mathbf{P}\{X \leq 2\} = \Phi(2) = 0.9772$.

(f) $\mathbf{P}\{X \geq -2\} = \mathbf{P}\{X \leq 2\} = \Phi(2) = 0.9772$.

(g) $\mathbf{P}\{-2 \leq X < 3\} = \mathbf{P}\{X < 3\} - \mathbf{P}\{X < -2\} = \mathbf{P}\{X \leq 3\} - \mathbf{P}\{X \leq -2\} = \Phi(3) - \Phi(-2) = (0.9987) - (1 - 0.9772) = 0.9987 - 0.0228 = 0.9759$.

(h) $\mathbf{P}\{-1 \leq X \leq 3\} = \mathbf{P}\{X < 3\} - \mathbf{P}\{X < -1\} = \mathbf{P}\{X \leq 3\} - \mathbf{P}\{X \leq -1\} = \Phi(3) - \Phi(-1) = 0.9987 - 0.1587 = 0.8400$.

6. (a) Let A_j , $j = 1, 2, 3$, denote the event that the lifetime of the j th TV lasts for at least two years. Therefore, since the TVs are selected at random and the TV lifetimes are independent, we find

$$\mathbf{P}\{A_1\} = \mathbf{P}\{A_2\} = \mathbf{P}\{A_3\} = \mathbf{P}\{X \geq 2\} = \int_2^{\infty} f(x) dx = \int_2^{\infty} \frac{1}{2} e^{-x/2} dx = -e^{-x/2} \Big|_2^{\infty} = \frac{1}{e}.$$

Hence, the probability that all three TVs last for at least two years is

$$\mathbf{P}\{A_1 \cap A_2 \cap A_3\} = \mathbf{P}\{A_1\} \mathbf{P}\{A_2\} \mathbf{P}\{A_3\} = \left(\frac{1}{e}\right)^3 = e^{-3}.$$

(b) Let B_j , $j = 1, 2, 3$, denote the event that the lifetime of the j th TV lasts for less than one year. Therefore, since the TVs are selected at random and the TV lifetimes are independent, we find

$$\mathbf{P}\{B_1\} = \mathbf{P}\{B_2\} = \mathbf{P}\{B_3\} = \mathbf{P}\{X < 1\} = \int_{-\infty}^1 f(x) dx = \int_0^1 \frac{1}{2} e^{-x/2} dx = -e^{-x/2} \Big|_0^1 = 1 - e^{-1/2}.$$

Hence, the probability that exactly one TV last for less than one year

$$\begin{aligned} & \mathbf{P}\{\text{exactly one TV lasts for less than one year}\} \\ &= \mathbf{P}\{B_1 \cap B_2^c \cap B_3^c \text{ or } B_1^c \cap B_2 \cap B_3^c \text{ or } B_1^c \cap B_2^c \cap B_3\} \\ &= \mathbf{P}\{B_1 \cap B_2^c \cap B_3^c\} + \mathbf{P}\{B_1^c \cap B_2 \cap B_3^c\} + \mathbf{P}\{B_1^c \cap B_2^c \cap B_3\} \\ &= \mathbf{P}\{B_1\} \mathbf{P}\{B_2^c\} \mathbf{P}\{B_3^c\} + \mathbf{P}\{B_1^c\} \mathbf{P}\{B_2\} \mathbf{P}\{B_3^c\} + \mathbf{P}\{B_1^c\} \mathbf{P}\{B_2^c\} \mathbf{P}\{B_3\} \\ &= (1 - e^{-1/2})(e^{-1/2})^2 + (1 - e^{-1/2})(e^{-1/2})^2 + (1 - e^{-1/2})(e^{-1/2})^2 \\ &= 3e^{-1}(1 - e^{-1/2}). \end{aligned}$$

7. (You may want to draw a tree diagram to help interpret the solution.) Let A be the event that a randomly selected Toyota vehicle is recalled. Let B_1 be the event that a randomly selected Toyota vehicle is a car, let B_2 be the event that a randomly selected Toyota vehicle is a truck, and let B_3 be the event that a randomly selected Toyota vehicle is a van. We are told that $\mathbf{P}\{B_1\} = 0.65$, $\mathbf{P}\{B_2\} = 0.20$, and $\mathbf{P}\{B_3\} = 0.15$. We are also told that $\mathbf{P}\{A|B_1\} = 0.10$, $\mathbf{P}\{A|B_2\} = 0.08$, and $\mathbf{P}\{A|B_3\} = 0.12$. We want to determine $\mathbf{P}\{B_2|A\}$. Using Bayes' Rule we find

$$\begin{aligned} \mathbf{P}\{B_2|A\} &= \frac{\mathbf{P}\{A|B_2\} \mathbf{P}\{B_2\}}{\mathbf{P}\{A\}} = \frac{\mathbf{P}\{A|B_2\} \mathbf{P}\{B_2\}}{\mathbf{P}\{A|B_1\} \mathbf{P}\{B_1\} + \mathbf{P}\{A|B_2\} \mathbf{P}\{B_2\} + \mathbf{P}\{A|B_3\} \mathbf{P}\{B_3\}} \\ &= \frac{(0.08)(0.20)}{(0.10)(0.65) + (0.08)(0.20) + (0.12)(0.15)} \\ &= \frac{16}{99} \doteq 0.161616. \end{aligned}$$