

ON THE WEAK LAW FOR RANDOMLY INDEXED PARTIAL SUMS FOR ARRAYS

DUG HUN HONG, SOO HAK SUNG, AND ANDREI I. VOLODIN

ABSTRACT. For randomly indexed sums of the form $\sum_{i=1}^{N_n}(X_{ni} - c_{ni})/b_n$, where $\{X_{ni}, i \geq 1, n \geq 1\}$ are random variables, $\{N_n, n \geq 1\}$ are positive integer-valued random variables, $\{c_{ni}, i \geq 1, n \geq 1\}$ are suitable conditional expectations and $\{b_n, n \geq 1\}$ are positive constants, we establish a general weak law of large numbers. Our result improves that of Hong [3].

Introduction

Let $\{X_{ni}, i \geq 1, n \geq 1\}$ be an array of random variables on the probability space (Ω, \mathcal{F}, P) and set $\mathcal{F}_{n,j} = \sigma\{X_{ni}, 1 \leq i \leq j\}, j \geq 1, n \geq 1$, and $\mathcal{F}_{n,0} = \{\emptyset, \Omega\}, n \geq 1$. Let $\{N_n, n \geq 1\}$ be a sequence of positive integer-valued random variables. In this paper, we establish a general weak law of large numbers(WLLN) of the form

$$(1) \quad \frac{\sum_{i=1}^{N_n}(X_{ni} - c_{ni})}{b_n} \rightarrow 0 \quad \text{in probability}$$

as $n \rightarrow \infty$, where $\{c_{ni}, i \geq 1, n \geq 1\}$ is an array of random variables and $\{b_n, n \geq 1\}$ is a sequence of positive constants. Hong [3] obtained a WLLN of the following form under the assumption that $N_n/b_n \rightarrow c$ in probability as $n \rightarrow \infty$, where c ($c > 0$) is constant and

$$(2) \quad \frac{\sum_{i=1}^{N_n}(X_{ni} - c_{ni})}{N_n} \rightarrow 0 \quad \text{in probability}$$

as $n \rightarrow \infty$. Note that (1) not only implies (2) under the condition that $N_n/b_n \rightarrow c$ ($c \neq 0$) in probability as $n \rightarrow \infty$ but also gives the following

Received December 20, 2000.

2000 Mathematics Subject Classification: 60F05, 60G42.

Key words and phrases: weak law of large numbers, convergence in probability, arrays, randomly indexed sums, martingale difference sequences.

generalization

$$\frac{\sum_{i=1}^{N_n} (X_{ni} - c_{ni})}{\phi(N_n)} \rightarrow 0 \quad \text{in probability}$$

as $n \rightarrow \infty$ under the condition $\phi(N_n)/b_n \rightarrow c$ ($c \neq 0$) in probability as $n \rightarrow \infty$, where $\phi : N \rightarrow R$ is a real function. When $\{N_n, n \geq 1\}$ is a sequence of positive integers, the WLLNs of the form (1) have been established by Gut [2], Hong and Oh [4], Kowalski and Rychlik [5], and Sung [6]. Our result is a generalization and improvement of Hong's result [3]. The proof owes much to those of earlier articles.

Main results

To prove the main results, we need the following lemma.

LEMMA 1. *Let $\{X_{ni}, i \geq 1, n \geq 1\}$ be an array of random variables, and $\{b_n, n \geq 1\}$ be a sequence of constants satisfying*

$$(3) \quad b_n/n = O(1).$$

Suppose that

$$(4) \quad \frac{1}{m} \sum_{i=1}^m jP(|X_{ni}| > b_j) \rightarrow 0$$

as $j \rightarrow \infty$ uniformly in n and m . Then

$$\frac{1}{m} \sum_{i=1}^m jP(|X_{ni}| > j) \rightarrow 0$$

as $j \rightarrow \infty$ uniformly in n and m .

PROOF. By (3), there exists a constant C such that $b_n \leq Cn$ for n sufficiently large. It follows that

$$\begin{aligned} \sum_{i=1}^m jP(|X_{ni}| > j) &= \sum_{i=1}^m jP(|X_{ni}| > C \frac{j}{C}) \\ &\leq \sum_{i=1}^m jP(|X_{ni}| > C \lceil \frac{j}{C} \rceil) \leq \sum_{i=1}^m jP(|X_{ni}| > b_{\lceil j/C \rceil}) \\ &= \sum_{i=1}^m \frac{j}{\lceil j/C \rceil} \lceil \frac{j}{C} \rceil P(|X_{ni}| > b_{\lceil j/C \rceil}), \end{aligned}$$

where $[a]$ denotes the integer part of a . If $j \geq 2C$, then

$$\frac{j}{[j/C]} \leq \frac{j}{j/C - 1} = \frac{jC}{j - C} \leq \frac{jC}{j/2} = 2C.$$

Thus we have for $j \geq 2C$

$$\frac{1}{m} \sum_{i=1}^m jP(|X_{ni}| > j) \leq \frac{2C}{m} \sum_{i=1}^m [j/C]P(|X_{ni}| > b_{[j/C]}),$$

and so the result is proved by (4). □

Now, we state and prove our main results.

THEOREM 2. *Let $\{X_{ni}, i \geq 1, n \geq 1\}$ be an array of random variables, and $\{b_n, n \geq 1\}$ be a sequence of constants satisfying (3) and $0 < b_n \rightarrow \infty$. Let $\{N_n, n \geq 1\}$ be a sequence of positive integer-valued random variables such that*

$$(5) \quad P(N_n > kb_n) = o(1)$$

for some positive integer k . Suppose that (4) holds. Then

$$\frac{\sum_{i=1}^{N_n} (X_{ni} - E(X_{ni}I(|X_{ni}| \leq b_n)|\mathcal{F}_{n,i-1}))}{b_n} \rightarrow 0 \quad \text{in probability}$$

as $n \rightarrow \infty$.

PROOF. Let $X'_{ni} = X_{ni}I(|X_{ni}| \leq b_n)$ for $i \geq 1, n \geq 1$. Then we have by (3), (4), and (5) that

$$\begin{aligned} P\left(\left|\sum_{i=1}^{N_n} X_{ni} - \sum_{i=1}^{N_n} X'_{ni}\right| > b_n\epsilon\right) &\leq P(N_n > kb_n) + P(\cup_{i=1}^{kb_n} (X_{ni} \neq X'_{ni})) \\ &\leq o(1) + \frac{kb_n}{n} \frac{1}{kb_n} \sum_{i=1}^{kb_n} nP(|X_{ni}| > b_n) = o(1). \end{aligned}$$

Thus

$$\frac{\sum_{i=1}^{N_n} X_{ni} - \sum_{i=1}^{N_n} X'_{ni}}{b_n} \rightarrow 0 \quad \text{in probability}$$

as $n \rightarrow \infty$, and so it suffices to show that

$$\frac{\sum_{i=1}^{N_n} (X'_{ni} - E(X'_{ni}|\mathcal{F}_{n,i-1}))}{b_n} \rightarrow 0 \quad \text{in probability}$$

as $n \rightarrow \infty$. For $n \geq 1$ and $m \geq 1$, denote:

$$B_m^n = \{|\sum_{i=1}^m (X'_{ni} - E(X'_{ni}|\mathcal{F}_{n,i-1}))| > b_n \epsilon\}, \quad D_n = \cup_{m=1}^{kb_n} B_m^n.$$

Then by (5)

$$P(B_{N_n}^n) \leq P(B_{N_n}^n, N_n \leq kb_n) + P(N_n > kb_n) \leq P(D_n) + o(1),$$

and hence it is sufficient to show that $P(D_n) = o(1)$. Since $\{X'_{ni} - E(X'_{ni}|\mathcal{F}_{n,i-1}), 1 \leq i \leq kb_n\}$ is a martingale difference sequence, the Hájek-Rényi inequality (see Theorem 7.4.8 in [1]) implies

$$\begin{aligned} P(D_n) &= P(\max_{1 \leq m \leq kb_n} |\sum_{i=1}^m (X'_{ni} - E(X'_{ni}|\mathcal{F}_{n,i-1}))| > b_n \epsilon) \\ &\leq \frac{1}{b_n^2 \epsilon^2} \sum_{i=1}^{kb_n} E(X'_{ni} - E(X'_{ni}|\mathcal{F}_{n,i-1}))^2 \leq \frac{1}{b_n^2 \epsilon^2} \sum_{i=1}^{kb_n} E|X'_{ni}|^2. \end{aligned}$$

Moreover,

$$\begin{aligned} \sum_{i=1}^{kb_n} E|X'_{ni}|^2 &= \sum_{i=1}^{kb_n} \sum_{j=1}^{b_n} E|X_{ni}|^2 I(j-1 < |X_{ni}| \leq j) \\ &\leq \sum_{i=1}^{kb_n} \sum_{j=1}^{b_n} j^2 P(j-1 < |X_{ni}| \leq j) \\ &= \sum_{i=1}^{kb_n} \sum_{j=1}^{b_n} j^2 [P(|X_{ni}| > j-1) - P(|X_{ni}| > j)] \\ &= \sum_{i=1}^{kb_n} [P(|X_{ni}| > 0) - b_n^2 P(|X_{ni}| > b_n) \\ &\quad + \sum_{j=1}^{b_n-1} ((j+1)^2 - j^2) P(|X_{ni}| > j)] \\ &\leq kb_n + \sum_{i=1}^{kb_n} \sum_{j=1}^{b_n-1} (2j+1) P(|X_{ni}| > j). \end{aligned}$$

By Lemma 1, we have

$$\lim_{j \rightarrow \infty} \sup_{n \geq 1} \frac{1}{kb_n} \sum_{i=1}^{kb_n} j P(|X_{ni}| > j) = 0.$$

Thus by the Toeplitz lemma we have

$$\begin{aligned} & \frac{1}{b_n^2} \sum_{i=1}^{kb_n} \sum_{j=1}^{b_n-1} (2j+1)P(|X_{ni}| > j) \\ &= \frac{k}{b_n} \sum_{j=1}^{b_n-1} \frac{2j+1}{j} \frac{1}{kb_n} \sum_{i=1}^{kb_n} jP(|X_{ni}| > j) \\ &= o(1), \end{aligned}$$

which implies $P(D_n) = o(1)$, since $b_n \rightarrow \infty$. □

COROLLARY 3. Let $\{X_{ni}\}$, $\{b_n\}$, and $\{N_n\}$ be as in Theorem 2 except that condition (5) in Theorem 2 is replaced by

$$(6) \quad N_n/b_n \rightarrow c \quad \text{in probability}$$

as $n \rightarrow \infty$, where $c \neq 0$. Then

$$\frac{\sum_{i=1}^{N_n} (X_{ni} - E(X_{ni}I(|X_{ni}| \leq b_n)|\mathcal{F}_{n,i-1}))}{N_n} \rightarrow 0 \quad \text{in probability}$$

as $n \rightarrow \infty$.

PROOF. Take a positive integer k such that $c + \epsilon \leq k$. It follows from (6) that

$$P(N_n > kb_n) \leq P(|\frac{N_n}{b_n} - c| > \epsilon) = o(1).$$

Thus by Theorem 2 we have

$$\begin{aligned} & \frac{\sum_{i=1}^{N_n} (X_{ni} - E(X_{ni}I(|X_{ni}| \leq b_n)|\mathcal{F}_{n,i-1}))}{N_n} \\ &= \frac{b_n}{N_n} \frac{\sum_{i=1}^{N_n} (X_{ni} - E(X_{ni}I(|X_{ni}| \leq b_n)|\mathcal{F}_{n,i-1}))}{b_n} \rightarrow 0 \quad \text{in probability} \end{aligned}$$

as $n \rightarrow \infty$. □

REMARK. Hong [3] proved Corollary 3 under the stronger conditions that

$$(7) \quad \frac{b_n}{n} \downarrow, \quad \frac{b_n^2}{n} \uparrow \infty, \quad \text{and} \quad \sum_{i=1}^n \frac{b_i^2}{i^2} = O(\frac{b_n^2}{n}).$$

Note that the first condition of (7) implies (3) and the second condition implies $b_n \rightarrow \infty$.

References

- [1] Y. S. Chow and H. Teicher, *Probability Theory: Independence, Interchangeability, Martingales*, Springer-Verlag, New York, 3rd ed., 1997.
- [2] A. Gut, *The weak law of large numbers for arrays*, Statist. Probab. Lett. **14** (1992), 49–52.
- [3] D. H. Hong, *On the weak law of large numbers for randomly indexed partial sums for arrays*, Statist. Probab. Lett. **28** (1996), 127–130.
- [4] D. H. Hong and K. S. Oh, *On the weak law of large numbers for arrays*, Statist. Probab. Lett. **22** (1995), 55–57.
- [5] P. Kowalski and Z. Rychlik, *On the weak law of large numbers for randomly indexed partial sums for arrays*, Ann. Univ. Mariae Curie-Sklodowska Sect. A **LI.1** (1997), 109–119.
- [6] S. H. Sung, *Weak law of large numbers for arrays*, Statist. Probab. Lett. **38** (1998), 101–105.

Dug Hun Hong
School of Mech. and Autom. Engin.
Catholic University of Taegu
Kyunbuk 713-702, South Korea
E-mail: dhhong@cuth.cataegu.ac.kr

Soo Hak Sung
Department of Applied Mathematics
Pai Chai University
Taejon 302-735, South Korea
E-mail: sungsh@www.paichai.ac.kr

Andrei I. Volodin
Department of Mathematics and Stat.
Regina University
Regina, Saskatchewan S4S 0A2, Canada
E-mail: volodin@math.uregina.ca