

© 2017 г. WANG X. J.\*, HU S. H.\*, VOLODIN A. I.\*\*

MOMENT INEQUALITIES FOR  $m$ -NOD RANDOM  
VARIABLES AND THEIR APPLICATIONS<sup>1)</sup>

Вводится понятие  $m$ -отрицательно ортант зависимых (сокращенно  $m$ -NOD) случайных величин и для них устанавливаются моментные неравенства, такие как неравенство Марцинкевича–Зигмунда и Розенталя. Как одно из применений моментных неравенств изучаются  $L_r$ - и почти наверное сходимости для  $m$ -NOD случайных величин при определенных условиях на равномерную интегрируемость. С другой стороны, устанавливается асимптотическое разложение обратных моментов для неотрицательных  $m$ -NOD случайных величин с конечными начальными моментами. Результаты статьи обобщают или улучшают некоторые известные результаты для независимых и некоторых классов зависимых последовательностей.

*Ключевые слова и фразы:*  $m$ -отрицательно ортант зависимые случайные величины;  $L_r$ -сходимость; обратные моменты; неравенства Марцинкевича–Зигмунда; неравенства Розенталя.

DOI: <https://doi.org/10.4213/tvp5123>

**1. Introduction.** It is well known that the Marcinkiewicz–Zygmund type inequality and Rosenthal type inequality play important role in probability limit theory and mathematical statistics, especially in establishing strong convergence, complete convergence, weak convergence, consistency and asymptotic normality in many stochastic models. There are many sequences of random variables satisfying the Marcinkiewicz–Zygmund type inequality or Rosenthal type inequality under some suitable conditions, such as independent sequence,  $\varphi$ -mixing sequence with the mixing coefficients satisfying certain conditions (see [36]),  $\rho$ -mixing sequence with the mixing coefficients satisfying certain conditions (see [20]),  $\tilde{\rho}$ -mixing sequence (see [32])

\*School of Mathematical Sciences, Anhui University, China; e-mail: wxjahdx@126.com

\*\*Department of Mathematics and Statistics, University of Regina, Regina, Canada.

<sup>1)</sup>Supported by the National Natural Science Foundation of China (11671012, 11501004, 11501005), the Natural Science Foundation of Anhui Province (1508085J06), and the Key Projects for Academic Talent of Anhui Province (gxbjZD2016005).

or [40]), negatively associated sequence (NA, in short, see [21]), negatively orthant dependent sequence (NOD, in short, see [2]), extended negatively dependent sequence (END, in short, see [22]), negatively superadditive dependent sequence (NSD, in short, see [11] or [39]), asymptotically almost negatively associated sequence with the mixing coefficients satisfying certain conditions (AANA, in short, see [49]),  $\rho^-$ -mixing sequence with the mixing coefficients satisfying certain conditions (see [34]), and so on.

The main purpose of the paper is to introduce a new concept of dependent structure —  $m$ -negatively orthant dependence ( $m$ -NOD, in short) and establish the Marcinkiewicz–Zygmund type inequality and Rosenthal type inequality for  $m$ -NOD random variables. In addition, we will give some applications of Marcinkiewicz–Zygmund type inequality and Rosenthal type inequality to  $L_r$  convergence, strong law of large numbers and the asymptotic approximation of inverse moments for nonnegative  $m$ -NOD random variables with finite first moments.

Firstly, let us recall the definition of negatively orthant dependent random variables which was introduced by Joav-Dev and Proschan [14] as follows.

*Definition 1.1.* A finite collection of random variables  $X_1, \dots, X_n$  is said to be *negatively orthant dependent* (NOD, in short) if both

$$\mathbf{P}(X_1 > x_1, \dots, X_n > x_n) \leq \prod_{i=1}^n \mathbf{P}(X_i > x_i)$$

and

$$\mathbf{P}(X_1 \leq x_1, \dots, X_n \leq x_n) \leq \prod_{i=1}^n \mathbf{P}(X_i \leq x_i)$$

hold for each  $n \geq 1$  and all real numbers  $x_1, \dots, x_n$ . An infinite sequence  $\{X_n, n \geq 1\}$  is said to be NOD if every finite subcollection is NOD.

An array  $\{X_{ni}, i \geq 1, n \geq 1\}$  of random variables is said to be rowwise NOD if for every  $n \geq 1$ ,  $\{X_{ni}, i \geq 1\}$  is a sequence of NOD random variables.

The class of NOD random variables is a very general dependent structure, which includes independent random variables and NA random variables as special cases. For more details about the probability inequalities, moment inequalities, or probability limit theory and applications, one can refer to [14], [4], [33], [31], [16], [2], [41], [37], [38], [42], [50], [19], [25], [45], and etc.

Inspired by the definition of NOD random variables, we introduce the concept of  $m$ -negatively orthant dependent random variables as follows.

*Definition 1.2.* Let  $m \geq 1$  be a fixed integer. A sequence  $\{X_n, n \geq 1\}$  of random variables is said to be  *$m$ -negatively orthant dependent* ( $m$ -NOD, in short) if for any  $n \geq 2$  and any  $i_1, \dots, i_n$  such that  $|i_k - i_j| \geq m$  for all  $1 \leq k \neq j \leq n$ , we have that  $X_{i_1}, \dots, X_{i_n}$  are negatively orthant dependent.

An array  $\{X_{ni}, i \geq 1, n \geq 1\}$  of random variables is said to be rowwise  $m$ -NOD if for every  $n \geq 1$ ,  $\{X_{ni}, i \geq 1\}$  is a sequence of  $m$ -NOD random variables.

When  $n = 2$ ,  $m$ -NOD reduces to  $m$ -pairwise NOD which was introduced by Anh in [1], and carefully studied by Wu and Rosalsky in [46]. When  $m = 1$ , the concept  $m$ -NOD random variables reduces to the so-called NOD random variables. Hence, the concept of  $m$ -NOD random variables is a natural extension from NOD random variables. Joav-Dev and Proschan in [14] pointed out that NA implies NOD, but NOD does not implies NA. Hu and Yang [12] or Hu [13] pointed out that NSD implies NOD. Hence, the class of  $m$ -NOD random variables includes independent random variables, NA random variables, NSD random variables, NOD random variables, and  $m$ -NA random variables (see [10]) as special cases. Studying the probability inequalities, moment inequalities, limiting behavior of  $m$ -NOD random variables and their applications in many stochastic models are of great interest.

The following lemmas for NOD random variables will be used in establishing the Marcinkiewicz–Zygmund type inequality and Rosenthal type inequality for  $m$ -NOD random variables.

**Lemma 1.1** (cf. [4]). *Let random variables  $X_1, \dots, X_n$  be NOD,  $f_1, \dots, f_n$  be all nondecreasing (or all nonincreasing) functions, then random variables  $f_1(X_1), \dots, f_n(X_n)$  are NOD.*

**Lemma 1.2** (cf. [2]). *Let  $\{X_n, n \geq 1\}$  be a sequence of NOD random variables with  $\mathbf{E}X_n = 0$  and  $\mathbf{E}|X_n|^p < \infty$  for some  $p \geq 1$  and every  $n \geq 1$ . Then there exist positive constants  $C_p$  and  $D_p$  depending only on  $p$  such that for every  $n \geq 1$ ,*

$$\mathbf{E} \left| \sum_{i=1}^n X_i \right|^p \leq C_p \sum_{i=1}^n \mathbf{E}|X_i|^p \quad \text{for } 1 \leq p \leq 2 \tag{1.1}$$

and

$$\mathbf{E} \left| \sum_{i=1}^n X_i \right|^p \leq D_p \left\{ \sum_{i=1}^n \mathbf{E}|X_i|^p + \left( \sum_{i=1}^n \mathbf{E}X_i^2 \right)^{p/2} \right\} \quad \text{for } p > 2. \tag{1.2}$$

Throughout the paper, let  $C$  denote a positive constant not depending on  $n$ , which may be different in various places;  $a_n = O(b_n)$  stands for  $a_n \leq Cb_n$ , where  $\{a_n, n \geq 1\}$  and  $\{b_n, n \geq 1\}$  are sequences of nonnegative real numbers. Denote  $\ln x = \ln \max(x, e)$ ,  $x^+ = x\mathbf{I}(x > 0)$ ,  $x^- = -x\mathbf{I}(x < 0)$ .

This work is organized as follows: the Marcinkiewicz–Zygmund type inequality and Rosenthal type inequality for  $m$ -NOD random variables are provided in Section 2. Some results on  $L_r$  convergence and strong law of large numbers for arrays of rowwise  $m$ -NOD random variables are established in

Section 3. The asymptotic approximation of inverse moments for nonnegative  $m$ -NOD random variables with finite first moments is investigated in Section 4.

**2. Marcinkiewicz–Zygmund type inequality and Rosenthal type inequality for  $m$ -NOD random variables.** In this section, we will establish the Marcinkiewicz–Zygmund type inequality and Rosenthal type inequality for  $m$ -NOD random variables, which can be applied to prove the strong convergence,  $L_r$  convergence, weak convergence, complete convergence, consistency and asymptotic normality in many stochastic models, and so on. To prove the main results, we need the following lemma, which will be used frequently throughout the paper.

**Lemma 2.1.** *Let  $\{X_n, n \geq 1\}$  be a sequence of  $m$ -NOD random variables. If  $\{f_n(\cdot), n \geq 1\}$  are all nondecreasing (or nonincreasing) functions, then random variables  $\{f_n(X_n), n \geq 1\}$  are  $m$ -NOD.*

This lemma can be obtained easily by the definition of  $m$ -NOD random variables and Lemma 1.1. So the details are omitted.

With Lemma 1.2 and Lemma 2.1 accounted for, we can establish the Marcinkiewicz–Zygmund type inequality and Rosenthal type inequality for  $m$ -NOD random variables as follows.

**Theorem 2.1.** *Let  $\{X_n, n \geq 1\}$  be a sequence of  $m$ -NOD random variables with  $\mathbf{E}X_n = 0$  and  $\mathbf{E}|X_n|^p < \infty$  for some  $p \geq 1$  and every  $n \geq 1$ . Then there exist positive constants  $C_{m,p}$  and  $D_{m,p}$  depending only on  $m$  and  $p$  such that for every  $n \geq m$ ,*

$$\mathbf{E} \left| \sum_{i=1}^n X_i \right|^p \leq \begin{cases} C_{m,p} \sum_{i=1}^n \mathbf{E}|X_i|^p, & \text{for } 1 \leq p \leq 2, \\ D_{m,p} \left[ \sum_{i=1}^n \mathbf{E}|X_i|^p + \left( \sum_{i=1}^n \mathbf{E}X_i^2 \right)^{p/2} \right], & \text{for } p > 2, \end{cases} \tag{2.1}$$

and

$$\mathbf{E} \left( \max_{1 \leq j \leq n} \left| \sum_{i=1}^j X_i \right|^p \right) \leq \begin{cases} C_{m,p} \ln^p n \sum_{i=1}^n \mathbf{E}|X_i|^p, & \text{for } 1 \leq p \leq 2, \\ D_{m,p} \ln^p n \left[ \sum_{i=1}^n \mathbf{E}|X_i|^p + \left( \sum_{i=1}^n \mathbf{E}X_i^2 \right)^{p/2} \right], & \text{for } p > 2. \end{cases} \tag{2.2}$$

*Proof.* From (2.1), we can see that (2.2) can be obtained immediately by a similar way of the process of Theorem 2.3.1 in [28]. So we only need to prove (2.1).

For fixed  $n \geq m$ , let  $r = \lceil n/m \rceil$ . Define

$$Y_i = \begin{cases} X_i, & 1 \leq i \leq n, \\ 0, & i > n. \end{cases}$$

Denote  $S'_{mr+j} = \sum_{i=0}^r Y_{mi+j}$  for  $j = 1, \dots, m$ . Noting that  $\sum_{i=1}^n X_i = \sum_{j=1}^m S'_{mr+j}$ , we have by  $C_r$ -inequality that

$$\mathbf{E} \left| \sum_{i=1}^n X_i \right|^p = \mathbf{E} \left| \sum_{j=1}^m S'_{mr+j} \right|^p \leq m^{p-1} \sum_{j=1}^m \mathbf{E} |S'_{mr+j}|^p. \tag{2.3}$$

By definition of  $m$ -NOD random variables, we see that  $Y_j, Y_{m+j}, \dots, Y_{mr+j}$  are NOD random variables for each  $j = 1, \dots, m$ .

For  $1 \leq p \leq 2$ , we have by (1.1) and (2.3) that for any  $n \geq m$ ,

$$\begin{aligned} \mathbf{E} \left| \sum_{i=1}^n X_i \right|^p &\leq m^{p-1} C_p \sum_{j=1}^m \sum_{i=0}^r \mathbf{E} |Y_{mi+j}|^p \\ &\leq m^p C_p \sum_{i=1}^n \mathbf{E} |X_i|^p \doteq C_{m,p} \sum_{i=1}^n \mathbf{E} |X_i|^p. \end{aligned} \tag{2.4}$$

For  $p > 2$ , we have by (1.2) and (2.3) that for any  $n \geq m$ ,

$$\begin{aligned} \mathbf{E} \left| \sum_{i=1}^n X_i \right|^p &\leq m^{p-1} D_p \sum_{j=1}^m \left[ \sum_{i=0}^r \mathbf{E} |Y_{mi+j}|^p + \left( \sum_{i=0}^r \mathbf{E} Y_{mi+j}^2 \right)^{p/2} \right] \\ &\leq m^p D_p \left[ \sum_{i=1}^n \mathbf{E} |X_i|^p + \left( \sum_{i=1}^n \mathbf{E} X_i^2 \right)^{p/2} \right] \\ &\doteq D_{m,p} \left[ \sum_{i=1}^n \mathbf{E} |X_i|^p + \left( \sum_{i=1}^n \mathbf{E} X_i^2 \right)^{p/2} \right]. \end{aligned} \tag{2.5}$$

The desired result (2.1) follows by (2.4) and (2.5) immediately. This completes the proof of the theorem.

*Remark 2.1.* Assume that (2.1) holds for any  $n \geq m$  and  $\sum_{i=1}^\infty X_i$  converges almost surely. Then

$$\mathbf{E} \left| \sum_{i=1}^\infty X_i \right|^p \leq \begin{cases} C_{m,p} \sum_{i=1}^\infty \mathbf{E} |X_i|^p, & \text{for } 1 \leq p \leq 2, \\ D_{m,p} \left[ \sum_{i=1}^\infty \mathbf{E} |X_i|^p + \left( \sum_{i=1}^\infty \mathbf{E} X_i^2 \right)^{p/2} \right], & \text{for } p > 2. \end{cases} \tag{2.6}$$

In fact, it follows by Fatou’s lemma that

$$\begin{aligned} \mathbf{E} \left| \sum_{i=1}^{\infty} X_i \right|^p &= \mathbf{E} \left| \liminf_{n \rightarrow \infty} \sum_{i=1}^n X_i \right|^p \leq \mathbf{E} \left( \liminf_{n \rightarrow \infty} \left| \sum_{i=1}^n X_i \right|^p \right) \\ &\leq \liminf_{n \rightarrow \infty} \mathbf{E} \left| \sum_{i=1}^n X_i \right|^p \leq \limsup_{n \rightarrow \infty} \mathbf{E} \left| \sum_{i=1}^n X_i \right|^p, \end{aligned} \tag{2.7}$$

which together with (2.1) yields (2.6).

*Remark 2.2.* Let  $\{a_n, n \geq 1\}$  be a sequence of real numbers. Under the conditions of Theorem 2.1, we have for  $n \geq m$  that

$$\mathbf{E} \left| \sum_{i=1}^n a_i X_i \right|^p \leq \begin{cases} 2^{p-1} C_{m,p} \sum_{i=1}^n \mathbf{E} |a_i X_i|^p, & \text{for } 1 \leq p \leq 2, \\ 2^p D_{m,p} \left[ \sum_{i=1}^n \mathbf{E} |a_i X_i|^p + \left( \sum_{i=1}^n \mathbf{E} a_i^2 X_i^2 \right)^{p/2} \right], & \text{for } p > 2. \end{cases} \tag{2.8}$$

Assume further that  $\sum_{i=1}^{\infty} a_i X_i$  converges almost surely, we have for  $n \geq m$  that

$$\mathbf{E} \left| \sum_{i=1}^{\infty} a_i X_i \right|^p \leq \begin{cases} 2^{p-1} C_{m,p} \sum_{i=1}^{\infty} \mathbf{E} |a_i X_i|^p, & \text{for } 1 \leq p \leq 2, \\ 2^p D_{m,p} \left[ \sum_{i=1}^{\infty} \mathbf{E} |a_i X_i|^p + \left( \sum_{i=1}^{\infty} \mathbf{E} a_i^2 X_i^2 \right)^{p/2} \right], & \text{for } p > 2. \end{cases} \tag{2.9}$$

Actually, for fixed  $n \geq m$ ,  $\{a_i^+ X_i, 1 \leq i \leq n\}$  and  $\{a_i^- X_i, 1 \leq i \leq n\}$  are both  $m$ -NOD random variables from Lemma 2.1. Noting that  $a_{ni} = a_{ni}^+ - a_{ni}^-$ , we have by  $C_r$ -inequality that

$$\mathbf{E} \left| \sum_{i=1}^n a_i X_i \right|^p \leq 2^{p-1} \mathbf{E} \left| \sum_{i=1}^n a_i^+ X_i \right|^p + 2^{p-1} \mathbf{E} \left| \sum_{i=1}^n a_i^- X_i \right|^p. \tag{2.10}$$

Note that  $|a_i|^p = (a_i^+)^p + (a_i^-)^p$ , the desired result (2.8) follows by (2.1) and (2.10) immediately.

Similar to the proof of (2.7), we can get (2.9) by (2.8) immediately.

**3.  $L_r$  convergence and strong convergence for  $m$ -NOD random variables.** In the previous section, we established the Marcinkiewicz–Zygmund type inequality and Rosenthal type inequality for  $m$ -NOD random variables. As one application of the moment inequalities for  $m$ -NOD random variables, we will study the  $L_r$  convergence and strong convergence for  $m$ -NOD random variables under some uniformly integrable conditions.

In what follows, let  $\{u_n, n \geq 1\}$  and  $\{v_n, n \geq 1\}$  be two sequences of integers (not necessary positive or finite) such that  $v_n > u_n$  for all  $n \geq 1$  and  $v_n - u_n \rightarrow \infty$  as  $n \rightarrow \infty$ . Let  $\{k_n, n \geq 1\}$  be a sequence of positive numbers such that  $k_n \rightarrow \infty$  as  $n \rightarrow \infty$  and  $\{h(n), n \geq 1\}$  be an increasing sequence of positive constants with  $h(n) \uparrow \infty$  as  $n \uparrow \infty$ .

**3.1.  $L_r$  convergence and weak law of large numbers.** The notion of  $h$ -integrability for an array of random variables concerning an array of constant weights was introduced by Cabrera and Volodin [17] as follows.

*Definition 3.1.* Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of random variables and let  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of constants with  $\sum_{i=u_n}^{v_n} |a_{ni}| \leq C$  for all  $n \in \mathbf{N}$  and some constant  $C > 0$ . The array  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is said to be  $h$ -integrable with respect to the array of constants  $\{a_{ni}\}$  if

$$\sup_{n \geq 1} \sum_{i=u_n}^{v_n} |a_{ni}| \mathbf{E}|X_{ni}| < \infty \quad \text{and} \quad \lim_{n \rightarrow \infty} \sum_{i=u_n}^{v_n} |a_{ni}| \mathbf{E}|X_{ni}| \mathbf{I}(|X_{ni}| > h(n)) = 0.$$

The main idea of the notion of  $h$ -integrability with respect to the array of constants  $\{a_{ni}\}$  is to deal with weighted sums of random variables. Sung et al. [29] introduced a new concept of integrability which deals with usual normed sums of random variables as follows.

*Definition 3.2.* Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of random variables and  $r > 0$ . The array  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is said to be  $h$ -integrable with exponent  $r$  if

$$\sup_{n \geq 1} \frac{1}{k_n} \sum_{i=u_n}^{v_n} \mathbf{E}|X_{ni}|^r < \infty \quad \text{and} \quad \lim_{n \rightarrow \infty} \frac{1}{k_n} \sum_{i=u_n}^{v_n} \mathbf{E}|X_{ni}|^r \mathbf{I}(|X_{ni}|^r > h(n)) = 0.$$

Under the conditions of  $h$ -integrability with exponent  $r$  and  $h$ -integrability with respect to the array of constants  $\{a_{ni}\}$ , Sung et al. [29] obtained the following Theorem A and Theorem B for arrays of rowwise NA random variables, respectively.

**Theorem A.** Let  $1 \leq r < 2$ . Suppose that  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is an array of rowwise NA random variables. Let  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of constants. Assume that the following conditions hold:

- (i)  $\{|X_{ni}|^r\}$  is  $h$ -integrable concerning the array  $\{|a_{ni}|^r\}$ , i.e.,

$$\sup_{n \geq 1} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|X_{ni}|^r < \infty$$

and

$$\lim_{n \rightarrow \infty} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|X_{ni}|^r \mathbf{I}(|X_{ni}|^r > h(n)) = 0;$$

(ii)  $h(n) \sup_{u_n \leq i \leq v_n} |a_{ni}| \rightarrow 0$  as  $n \rightarrow \infty$ .

Then

$$\sum_{i=u_n}^{v_n} a_{ni}(X_{ni} - \mathbf{E}X_{ni}) \rightarrow 0$$

in  $L_r$  and, hence, in probability as  $n \rightarrow \infty$ .

**Theorem B.** Suppose that  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is an array of rowwise NA  $h$ -integrable with exponent  $1 \leq r < 2$  random variables,  $k_n \rightarrow \infty$ ,  $h(n) \uparrow \infty$ , and  $h(n)/k_n \rightarrow 0$  as  $n \rightarrow \infty$ . Then

$$\frac{\sum_{i=u_n}^{v_n} (X_{ni} - \mathbf{E}X_{ni})}{k_n^{1/r}} \rightarrow 0$$

in  $L_r$  and, hence, in probability as  $n \rightarrow \infty$ .

Inspired by the concept of  $h$ -integrability with exponent  $r$ , Wang and Hu [35] introduced a new and weaker concept of uniform integrability as follows.

*Definition 3.3.* Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of random variables and  $r > 0$ . The array  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is said to be *residually  $h$ -integrable* ( *$R$ - $h$ -integrable*, in short) with exponent  $r$  if

$$\sup_{n \geq 1} \frac{1}{k_n} \sum_{i=u_n}^{v_n} \mathbf{E}|X_{ni}|^r < \infty$$

and

$$\lim_{n \rightarrow \infty} \frac{1}{k_n} \sum_{i=u_n}^{v_n} \mathbf{E}(|X_{ni}| - h^{1/r}(n))^r \mathbf{I}(|X_{ni}|^r > h(n)) = 0.$$

Under the assumption of  $R$ - $h$ -integrability with exponent  $r$ , Wang and Hu [35] established some weak laws of large numbers for arrays of dependent random variables. Note that

$$(|X_{ni}| - h^{1/r}(n))^r \mathbf{I}(|X_{ni}|^r > h(n)) \leq |X_{ni}|^r \mathbf{I}(|X_{ni}|^r > h(n)),$$

hence, the concept of  $R$ - $h$ -integrability with exponent  $r$  is weaker than  $h$ -integrability with exponent  $r$ .

Just as  $h$ -integrability with exponent  $r$ , the main idea of the notion of  $R$ - $h$ -integrability with exponent  $r$  is used to deal with usual normed sums of random variables. We now introduce a new and weaker concept of integrability which deals with weighted sums of random variables.

*Definition 3.4.* Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of random variables and let  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of constants. Let  $r > 0$ . The array  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is said to be  *$R$ - $h$ -integrable with exponent  $r$  concerning the array  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$*  if

$$\sup_{n \geq 1} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|X_{ni}|^r < \infty$$



and

$$\lim_{n \rightarrow \infty} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}(|X_{ni}| - h^{1/r}(n))^r \mathbf{I}(|X_{ni}|^r > h(n)) = 0.$$

When  $r = 1$ , the notion of  $R$ - $h$ -integrability with exponent  $r$  concerning the array of constants  $\{a_{ni}\}$  reduces to the so-called  $R$ - $h$ -integrability concerning the array of constants  $\{a_{ni}\}$ . For more details about the  $L_r$  convergence for weighted sums of random variables based on  $R$ - $h$ -integrability, one can refer to [48], [25], and so on.

The main purpose of this section is to generalize and improve the results of Theorem A and Theorem B for arrays of rowwise NA random variables to the case of arrays of rowwise  $m$ -NOD random variables under some weaker conditions. In addition, we will study the  $L_r$  convergence and weak law of large numbers for a class of random variables under the condition of  $h$ -integrability with exponent  $1 \leq r < 2$ , which generalize the corresponding ones of [44] and [30], and improve the corresponding one of [6]. The key techniques used here are the Marcinkiewicz–Zygmund type inequality and the truncated method.

Our main results on  $L_r$  convergence and weak law of large numbers for arrays of rowwise  $m$ -NOD are as follows. The first one is based on the condition of  $R$ - $h$ -integrability with exponent  $1 \leq r < 2$  concerning the array of constants  $\{a_{ni}\}$ .

**Theorem 3.1.** *Let  $1 \leq r < 2$ . Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of rowwise  $m$ -NOD random variables and let  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of constants. Assume that the following conditions hold:*

(i)  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is  $R$ - $h$ -integrable with exponent  $r$  concerning the array of constants  $\{a_{ni}\}$ , i.e.,

$$\sup_{n \geq 1} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|X_{ni}|^r < \infty$$

and

$$\lim_{n \rightarrow \infty} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}(|X_{ni}| - h^{1/r}(n))^r \mathbf{I}(|X_{ni}|^r > h(n)) = 0;$$

(ii)  $h(n) \sup_{u_n \leq i \leq v_n} |a_{ni}|^r \rightarrow 0$  as  $n \rightarrow \infty$ .

Then

$$\sum_{i=u_n}^{v_n} a_{ni}(X_{ni} - \mathbf{E}X_{ni}) \rightarrow 0$$

in  $L_r$  and, hence, in probability as  $n \rightarrow \infty$ .

*Proof.* Since  $a_{ni} = a_{ni}^+ - a_{ni}^-$ , without loss of generality, we assume that  $a_{ni} \geq 0$ . For fixed  $n \geq 1$ , denote for  $u_n \leq i \leq v_n$  that

$$Y_{ni} = -h^{1/r}(n)\mathbf{I}(X_{ni} < -h^{1/r}(n)) + X_{ni}\mathbf{I}(|X_{ni}| \leq h^{1/r}(n))$$

$$\begin{aligned}
 &+ h^{1/r}(n)\mathbf{I}(X_{ni} > h^{1/r}(n)), \\
 Z_{ni} &= X_{ni} - Y_{ni} = (X_{ni} + h^{1/r}(n))\mathbf{I}(X_{ni} < -h^{1/r}(n)) \\
 &+ (X_{ni} - h^{1/r}(n))\mathbf{I}(X_{ni} > h^{1/r}(n)), \\
 S_n &= \sum_{i=u_n}^{v_n} a_{ni}(Y_{ni} - \mathbf{E}Y_{ni}), \quad T_n = \sum_{i=u_n}^{v_n} a_{ni}(Z_{ni} - \mathbf{E}Z_{ni}).
 \end{aligned}$$

Note that

$$\sum_{i=u_n}^{v_n} a_{ni}(X_{ni} - \mathbf{E}X_{ni}) = S_n + T_n, \quad n \geq 1,$$

we have by  $C_r$ -inequality that

$$\mathbf{E} \left| \sum_{i=u_n}^{v_n} a_{ni}(X_{ni} - \mathbf{E}X_{ni}) \right|^r \leq C\mathbf{E}|S_n|^r + C\mathbf{E}|T_n|^r.$$

To prove  $\sum_{i=u_n}^{v_n} a_{ni}(X_{ni} - \mathbf{E}X_{ni}) \rightarrow 0$  in  $L_r$ , we only need to show  $\mathbf{E}|S_n|^r \rightarrow 0$  and  $\mathbf{E}|T_n|^r \rightarrow 0$  as  $n \rightarrow \infty$ , where  $1 \leq r < 2$ .

Firstly, we will show that  $\mathbf{E}|S_n|^r \rightarrow 0$  as  $n \rightarrow \infty$ . Note that  $1 \leq r < 2$ , it suffices to show  $\mathbf{E}S_n^2 \rightarrow 0$  as  $n \rightarrow \infty$ .

For fixed  $n \geq 1$ , it follows by Lemma 2.1 that  $\{a_{ni}(Y_{ni} - \mathbf{E}Y_{ni}), u_n \leq i \leq v_n\}$  are  $m$ -NOD random variables. Note that  $|Y_{ni}| = \min\{|X_{ni}|, h^{1/r}(n)\}$ , we have by Theorem 2.1 and Remark 2.1 that

$$\begin{aligned}
 \mathbf{E}S_n^2 &= \mathbf{E} \left| \sum_{i=u_n}^{v_n} a_{ni}(Y_{ni} - \mathbf{E}Y_{ni}) \right|^2 \leq C \sum_{i=u_n}^{v_n} a_{ni}^2 \mathbf{E}Y_{ni}^2 \\
 &\leq Ch^{(2-r)/r}(n) \sup_{n \geq 1} |a_{ni}|^{2-r} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|Y_{ni}|^r \\
 &\leq C \left[ h(n) \sup_{u_n \leq i \leq v_n} |a_{ni}|^r \right]^{(2-r)/r} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|X_{ni}|^r \rightarrow 0 \text{ as } n \rightarrow \infty,
 \end{aligned}$$

which implies that  $\mathbf{E}S_n^2 \rightarrow 0$  as  $n \rightarrow \infty$  and, thus,  $\mathbf{E}|S_n|^r \rightarrow 0$  as  $n \rightarrow \infty$ .

Further, we will show that  $\mathbf{E}|T_n|^r \rightarrow 0$  as  $n \rightarrow \infty$ .

For fixed  $n \geq 1$ , it follows by Lemma 2.1 again that  $\{a_{ni}(Z_{ni} - \mathbf{E}Z_{ni}), u_n \leq i \leq v_n\}$  are  $m$ -NOD random variables. Note that

$$|Z_{ni}| = (|X_{ni}| - h^{1/r}(n))\mathbf{I}(|X_{ni}| > h^{1/r}(n)),$$

we have by Theorem 2.1 and Remark 2.1 again that

$$\begin{aligned}
 \mathbf{E}|T_n|^r &= \mathbf{E} \left| \sum_{i=u_n}^{v_n} a_{ni}(Z_{ni} - \mathbf{E}Z_{ni}) \right|^r \leq C \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|Z_{ni}|^r \\
 &= C \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}(|X_{ni}| - h^{1/r}(n))^r \mathbf{I}(|X_{ni}| > h(n)) \rightarrow 0 \text{ as } n \rightarrow \infty,
 \end{aligned}$$

which implies that  $\mathbf{E}|T_n|^r \rightarrow 0$  as  $n \rightarrow \infty$ . This completes the proof of the theorem.

If we take  $a_{ni} = k_n^{-1/r}$  for  $u_n \leq i \leq v_n$  and  $n \geq 1$ , then we can get the following result on  $L_r$  convergence and weak law of large numbers for arrays of rowwise  $m$ -NOD  $R$ - $h$ -integrable with exponent  $1 \leq r < 2$  random variables.

**Corollary 3.1.** *Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of rowwise  $m$ -NOD  $R$ - $h$ -integrable random variables with exponent  $1 \leq r < 2, k_n \rightarrow \infty, h(n) \uparrow \infty$ , and  $h(n)/k_n \rightarrow 0$  as  $n \rightarrow \infty$ . Then*

$$\frac{1}{k_n^{1/r}} \sum_{i=u_n}^{v_n} (X_{ni} - \mathbf{E}X_{ni}) \rightarrow 0$$

in  $L_r$  and, hence, in probability as  $n \rightarrow \infty$ .

*Remark 3.1.* Note that

$$(|X_{ni}| - h^{1/r}(n))^r \mathbf{I}(|X_{ni}|^r > h(n)) \leq |X_{ni}|^r \mathbf{I}(|X_{ni}|^r > h(n)),$$

and  $h(n) \sup_{u_n \leq i \leq v_n} |a_{ni}| \rightarrow 0$  implies  $h(n) \sup_{u_n \leq i \leq v_n} |a_{ni}|^r \rightarrow 0$  (here,  $1 \leq r < 2$  and  $h(n) \uparrow \infty$  as  $n \rightarrow \infty$ ), which imply that the conditions of Theorem 3.1 are weaker than those of Theorem A. Hence, the result of Theorem 3.1 generalizes and improves the corresponding one of Theorem A.

*Remark 3.2.* Since the concept of  $R$ - $h$ -integrability with exponent  $r$  is weaker than  $h$ -integrability with exponent  $r$  and  $m$ -NOD is weaker than NA, the result of Corollary 3.1 generalizes and improves the corresponding one of Theorem B.

Further, we will establish the  $L_r$ -convergence and weak law of large numbers for a class of random variables satisfying the Marcinkiewicz–Zygmund inequality with exponent 2, which includes  $m$ -NOD as a special case. The main ideas are inspired by [6] and [30].

We say that a sequence  $\{X_n, n \geq 1\}$  of random variables satisfies the Marcinkiewicz–Zygmund inequality with exponent 2, if for all  $n \geq 1$ ,

$$\mathbf{E} \left| \sum_{i=1}^n X_i \right|^2 \leq C \sum_{i=1}^n \mathbf{E}|X_i|^2,$$

where  $C$  is a positive constant not depending on  $n$ .

We say that an array  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  of random variables satisfies the Marcinkiewicz–Zygmund inequality with exponent 2, if for all  $n \geq 1$ ,

$$\mathbf{E} \left| \sum_{i=u_n}^{v_n} X_{ni} \right|^2 \leq C \sum_{i=u_n}^{v_n} \mathbf{E}|X_{ni}|^2,$$

where  $C$  is a positive constant not depending on  $n$ .

*Remark 3.3.* There are many sequences of mean zero random variables satisfying the Marcinkiewicz–Zygmund inequality with exponent 2, such as independent sequence, martingale difference sequence,  $\varphi$ -mixing sequence with the mixing coefficients satisfying certain conditions (see [36]),  $\rho$ -mixing sequence with the mixing coefficients satisfying certain conditions (see [20]),  $\tilde{\rho}$ -mixing sequence (see [32]), NA sequence (see [21]), NOD sequence (see [2]), END sequence (see [22]), NSD sequence (see [11] or [39]), AANA sequence with the mixing coefficients satisfying certain conditions (see [49]),  $\rho^-$ -mixing sequence with the mixing coefficients satisfying certain conditions (see [34]), pairwise negatively quadrant dependent sequence (PNQD, in short, see [17]),  $m$ -NOD sequence (see Theorem 2.1 in the paper), linearly negative quadrant dependent sequence (LNQD, in short, [51]), and so on.

Our main result on  $L_r$  convergence and weak law of large numbers for a class of random variables satisfying the Marcinkiewicz–Zygmund inequality with exponent 2 is as follows.

**Theorem 3.2.** *Let  $1 \leq r < 2$ . Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of random variables and let  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of constants. Assume that the following conditions hold:*

- (i)  $\sup_{n \geq 1} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|X_{ni}|^r < \infty$ ,
- (ii) for any  $\varepsilon > 0$ ,

$$\lim_{n \rightarrow \infty} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|X_{ni}|^r \mathbf{I}(|X_{ni}|^r > \varepsilon) = 0,$$

(iii) for any  $t > 0$ , the array  $\{Y_{ni} - \mathbf{E}Y_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  satisfies the Marcinkiewicz–Zygmund inequality with exponent 2, where

$$Y_{ni} = a_{ni}X_{ni}\mathbf{I}(|a_{ni}X_{ni}| \leq t^{1/r})$$

or

$$Y_{ni} = -t^{1/r}\mathbf{I}(a_{ni}X_{ni} < -t^{1/r}) + a_{ni}X_{ni}\mathbf{I}(|a_{ni}X_{ni}| \leq t^{1/r}) + t^{1/r}\mathbf{I}(a_{ni}X_{ni} > t^{1/r}).$$

Then

$$\sum_{i=u_n}^{v_n} a_{ni}(X_{ni} - \mathbf{E}X_{ni}) \rightarrow 0$$

in  $L_r$  and, hence, in probability as  $n \rightarrow \infty$ .

*Proof.* The proof is similar to that of [30]. So the details are omitted.

With Theorem 3.2 in hand and similar to the proof of Corollary 2.1 in [30], we can get the following corollary.

**Corollary 3.2.** *Let  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of constants satisfying  $k_n \doteq 1/\sup_{u_n \leq i \leq v_n} |a_{ni}|^r \rightarrow \infty$ ,  $0 < h(n) \uparrow \infty$ , and  $h(n)/k_n \rightarrow 0$*

as  $n \rightarrow \infty$ . Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of  $h$ -integrable random variables with exponent  $1 \leq r < 2$ . Assume further that the condition (iii) in Theorem 3.2 holds. Then

$$\sum_{i=u_n}^{v_n} a_{ni}(X_{ni} - \mathbf{E}X_{ni}) \rightarrow 0$$

in  $L_r$  and, hence, in probability as  $n \rightarrow \infty$ .

Taking  $a_{ni} = k_n^{-1/r}$  for  $u_n \leq i \leq v_n$  and  $n \geq 1$  in Corollary 3.2, we can get the following corollary immediately.

**Corollary 3.3.** Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of  $h$ -integrable with exponent  $1 \leq r < 2$  random variables,  $k_n \rightarrow \infty, 0 < h(n) \uparrow \infty$ , and  $h(n)/k_n \rightarrow 0$  as  $n \rightarrow \infty$ . Assume further that the condition (iii) in Theorem 3.2 holds, where  $a_{ni} = k_n^{-1/r}$  for  $u_n \leq i \leq v_n$  and  $n \geq 1$ . Then

$$\frac{\sum_{i=u_n}^{v_n} (X_{ni} - \mathbf{E}X_{ni})}{k_n^{1/r}} \rightarrow 0$$

in  $L_r$  and, hence, in probability as  $n \rightarrow \infty$ .

*Remark 3.4.* We have pointed out that PNQD sequence satisfies the Marcinkiewicz–Zygmund inequality with exponent 2 in Remark 3.3. Hence, the results of Theorem 3.2 and Corollaries 3.2 and 3.3 in the paper generalize the corresponding ones of Theorem 2.1 and Corollaries 2.1 and 2.2 for PNQD random variables in [30], respectively. In addition, note that LNQD implies PNQD (see [30]), hence, our results of Corollary 3.2 and Corollary 3.3 generalize the corresponding ones of Theorem 3.1 and Corollary 3.1 for LNQD random variables in [44], respectively.

*Remark 3.5.* Under the conditions of Corollary 3.2, Chen et al. in [6] established the  $L_1$  convergence and weak law of large numbers for arrays of rowwise  $h$ -integrable with exponent  $r = 1$  random variables satisfying the Marcinkiewicz–Zygmund inequality with exponent 2. Here, we established the  $L_r$  convergence and weak law of large numbers for arrays of rowwise  $h$ -integrable with exponent  $1 \leq r < 2$  random variables satisfying the Marcinkiewicz–Zygmund inequality with exponent 2. In addition, the condition « $k_n \doteq 1/\sup_{u_n \leq i \leq v_n} |a_{ni}|^r \rightarrow \infty, 0 < h(n) \uparrow \infty$  and  $h(n)/k_n \rightarrow 0$  as  $n \rightarrow \infty$ » in Corollary 3.2 in the paper is weaker than « $k_n \doteq 1/\sup_{u_n \leq i \leq v_n} |a_{ni}| \rightarrow \infty, 0 < h(n) \uparrow \infty$  and  $h(n)/k_n \rightarrow 0$  as  $n \rightarrow \infty$ » in Theorem 1 in [6]. Hence, our results of Theorem 3.2 and Corollary 3.2 generalize and improve the corresponding one of Theorem 1 in [6].

**3.2. Strong convergence.** In Section 3.1, we studied the  $L_r$  convergence and weak law of large numbers for arrays of rowwise  $m$ -NOD random variables under some uniformly integrable conditions. In order to establish the strong version of Theorem 3.1, we introduce the concept of strongly residual  $h$ -integrability with exponent  $r$  concerning the array of constants  $\{a_{ni}\}$  as follows.

*Definition 3.5.* Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of random variables and let  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of constants. Let  $r > 0$ . The array  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is said to be *strongly residually  $h$ -integrable* (*SR- $h$ -integrable*, for short) with exponent  $r$  concerning the array  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  if

$$\sup_{n \geq 1} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|X_{ni}|^r < \infty$$

and

$$\sum_{n=1}^{\infty} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}(|X_{ni}| - h^{1/r}(n))^r \mathbf{I}(|X_{ni}|^r > h(n)) < \infty.$$

When  $r = 1$ , the preceding definition reduces to the concept of *SR- $h$ -integrability* concerning the array  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$ , which was introduced by Ordóñez Cabrera et al. [18].

The main idea of the notion of *SR- $h$ -integrability* with exponent  $r$  concerning the array  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is to deal with weighted sums of random variables. We introduce a new concept of integrability which deals with usual normed sums of random variables as follows.

*Definition 3.6.* Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of random variables and  $r > 0$ . The array  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is said to be *SR- $h$ -integrable with exponent  $r$*  if

$$\sup_{n \geq 1} \frac{1}{k_n} \sum_{i=u_n}^{v_n} \mathbf{E}|X_{ni}|^r < \infty$$

and

$$\sum_{n=1}^{\infty} \frac{1}{k_n} \sum_{i=u_n}^{v_n} \mathbf{E}(|X_{ni}| - h^{1/r}(n))^r \mathbf{I}(|X_{ni}|^r > h(n)) < \infty.$$

*Remark 3.6.* It is easily seen that the concept of *SR- $h$ -integrability* with exponent  $r$  is stronger than the concept of *R- $h$ -integrability* with exponent  $r$ .

Our main result on strong convergence for weighted sums of arrays of rowwise *m-NOD* random variables under some uniformly integrable conditions is as follows.

**Theorem 3.3.** *Let  $1 \leq r < 2$ . Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of rowwise *m-NOD* random variables and let  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of constants. Assume that the following conditions hold:*

- (i)  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  is *SR- $h$ -integrable with exponent  $r$*  concerning the array  $\{a_{ni}, u_n \leq i \leq v_n, n \geq 1\}$ ;
- (ii)

$$\sum_{n=1}^{\infty} \left( h(n) \sup_{u_n \leq i \leq v_n} |a_{ni}|^r \right)^{(2-r)/r} < \infty.$$

Then  $\sum_{i=u_n}^{v_n} a_{ni}(X_{ni} - \mathbf{E}X_{ni}) \rightarrow 0$  a.s. as  $n \rightarrow \infty$ .

*Proof.* We use the same notation as those in Theorem 3.1. To prove  $\sum_{i=u_n}^{v_n} a_{ni}(X_{ni} - \mathbf{E}X_{ni}) \rightarrow 0$  a.s. as  $n \rightarrow \infty$ , it suffices to show that

$$S_n \doteq \sum_{i=u_n}^{v_n} a_{ni}(Y_{ni} - \mathbf{E}Y_{ni}) \rightarrow 0 \quad \text{a.s. as } n \rightarrow \infty \tag{3.1}$$

and

$$T_n \doteq \sum_{i=u_n}^{v_n} a_{ni}(Z_{ni} - \mathbf{E}Z_{ni}) \rightarrow 0 \quad \text{a.s. as } n \rightarrow \infty. \tag{3.2}$$

Firstly, we will prove (3.1). Note that  $|Y_{ni}| = \min\{|X_{ni}|, h^{1/r}(n)\}$ , we have by Markov’s inequality, Theorem 2.1 (or Remark 2.1), Jensen’s inequality, and conditions (i), (ii) that for any  $\varepsilon > 0$ ,

$$\begin{aligned} \sum_{n=1}^{\infty} \mathbf{P}(|S_n| > \varepsilon) &\leq \frac{1}{\varepsilon^2} \sum_{n=1}^{\infty} \mathbf{E} \left| \sum_{i=u_n}^{v_n} a_{ni}(Y_{ni} - \mathbf{E}Y_{ni}) \right|^2 \leq C \sum_{n=1}^{\infty} \sum_{i=u_n}^{v_n} a_{ni}^2 \mathbf{E}Y_{ni}^2 \\ &\leq C \sum_{n=1}^{\infty} h^{(2-r)/r}(n) \sup_{u_n \leq i \leq v_n} |a_{ni}|^{2-r} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|Y_{ni}|^r \\ &\leq C \sum_{n=1}^{\infty} \left[ h(n) \sup_{u_n \leq i \leq v_n} |a_{ni}|^r \right]^{(2-r)/r} \left( \sup_{n \geq 1} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|X_{ni}|^r \right) < \infty, \end{aligned}$$

which implies (3.1) by Borel–Cantelli lemma.

In the following, we will prove (3.2). Note that

$$|Z_{ni}| = (|X_{ni}| - h^{1/r}(n)) \mathbf{I}(|X_{ni}|^r > h(n)),$$

we have by Markov’s inequality, Theorem 2.1 (or Remark 2.1), Jensen’s inequality and condition (i) that for any  $\varepsilon > 0$ ,

$$\begin{aligned} \sum_{n=1}^{\infty} \mathbf{P}(|T_n| > \varepsilon) &\leq \frac{1}{\varepsilon^r} \sum_{n=1}^{\infty} \mathbf{E} \left| \sum_{i=u_n}^{v_n} a_{ni}(Z_{ni} - \mathbf{E}Z_{ni}) \right|^r \leq C \sum_{n=1}^{\infty} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}|Z_{ni}|^r \\ &= C \sum_{n=1}^{\infty} \sum_{i=u_n}^{v_n} |a_{ni}|^r \mathbf{E}(|X_{ni}| - h^{1/r}(n))^r \mathbf{I}(|X_{ni}|^r > h(n)) < \infty, \end{aligned}$$

which implies (3.2) by Borel–Cantelli lemma. This completes the proof of the theorem.

If we take  $a_{ni} = k_n^{-1/r}$  for  $u_n \leq i \leq v_n$  and  $n \geq 1$  in Theorem 3.3, then we can get the following result on strong convergence for arrays of rowwise  $m$ -NOD  $SR$ - $h$ -integrable with exponent  $1 \leq r < 2$  random variables.

**Corollary 3.4.** Let  $\{X_{ni}, u_n \leq i \leq v_n, n \geq 1\}$  be an array of row-wise  $m$ -NOD SR- $h$ -integrable random variables with exponent  $1 \leq r < 2$ ,  $k_n \rightarrow \infty$ ,  $h(n) \uparrow \infty$ , and

$$\sum_{n=1}^{\infty} \left( \frac{h(n)}{k_n} \right)^{(2-r)/r} < \infty.$$

Then

$$\frac{1}{k_n^{1/r}} \sum_{i=u_n}^{v_n} (X_{ni} - \mathbf{E}X_{ni}) \rightarrow 0 \quad \text{a.s.}$$

as  $n \rightarrow \infty$ .

**4. On the asymptotic approximation of inverse moment for nonnegative  $m$ -NOD random variables.** As one application of the moment inequalities for  $m$ -NOD random variables, Section 3 deals with the  $L_r$  convergence and strong convergence for  $m$ -NOD random variables under some uniformly integrable conditions. As another application of the moment inequalities for  $m$ -NOD random variables, we will study the asymptotic approximation of inverse moments for nonnegative  $m$ -NOD random variables with finite first moments in this section.

Let  $Z_1, Z_2, \dots$  be a sequence of nonnegative random variables with finite second moments. Denote

$$X_n = \frac{1}{B_n} \sum_{i=1}^n Z_i, \quad B_n^2 = \sum_{i=1}^n D(Z_i). \quad (4.1)$$

Under some suitable conditions, the inverse moment can be approximated by the inverse of the moment in the following way:

$$\mathbf{E}(a + X_n)^{-\alpha} \sim (a + \mathbf{E}X_n)^{-\alpha}, \quad (4.2)$$

where  $a > 0$  and  $\alpha > 0$  are arbitrary real numbers. Here and in what follows, for two positive sequences  $\{c_n, n \geq 1\}$  and  $\{d_n, n \geq 1\}$ , we write  $c_n \sim d_n$  and  $c_n = o(d_n)$  if  $\lim_{n \rightarrow \infty} c_n d_n^{-1} = 1$ ,  $\lim_{n \rightarrow \infty} c_n d_n^{-1} = 0$ . The left-hand side of (4.2) is the inverse moment and the right-hand side is the inverse of the moment. Usually, the inverse of the moment is much easier to compute than the inverse moment. So in many practical applications, such as evaluating risks of estimators and powers of tests, reliability, life testing, insurance, and financial mathematics, complex systems, and so on, we often take the inverse of the moment instead of the inverse moment. Up to now, many authors studied the asymptotic approximation of inverse moment and found many interesting results. For the details about the inverse moment, one can refer to [5], [7], [8], [15], [43], [37], [26], [24], [47] [9], [27], and etc.



For  $n \geq 1$ , denote

$$\tilde{X}_n = \sum_{i=1}^n Z_i, \quad \tilde{\mu}_n = \mathbf{E}\tilde{X}_n \tag{4.3}$$

and

$$\mu_{n,s} = \sum_{i=1}^n \mathbf{E}Z_i \mathbf{I}(Z_i \leq \mu_n^s / \sqrt{n}) \quad \text{for some } 0 < s < 1.$$

Based on notation above, Shi et al. [26] obtained the following Theorem C and Theorem D.

**Theorem C.** *Let  $\{Z_n, n \geq 1\}$  be a sequence of independent, nonnegative, and nondegenerated random variables. Assume that the following conditions hold:*

- (H<sub>1</sub>)  $\tilde{\mu}_n \rightarrow \infty$  as  $n \rightarrow \infty$ ;
- (H<sub>2</sub>)  $\tilde{\mu}_n \sim \mu_{n,s}$  for some  $0 < s < 1$ .

Then

$$\mathbf{E}(a + \tilde{X}_n)^{-\alpha} \sim (a + \mathbf{E}\tilde{X}_n)^{-\alpha} \tag{4.4}$$

holds for all real constants  $a > 0$  and  $\alpha > 0$ .

**Theorem D.** *Let the conditions of Theorem C hold. In addition, suppose that there exists a function  $f(x), x \geq 0$ , satisfying the following conditions:*

- (H<sub>3</sub>) *there exists a  $c_1 > 0$  such that  $f(x) > c_1$  for  $x \geq 0$ ;*
- (H<sub>4</sub>) *there exist  $k > 0$  and  $c_2 > 0$  such that  $f(x)/x^k \rightarrow c_2$  as  $x \rightarrow \infty$ ;*
- (H<sub>5</sub>)  *$1/f(x)$  is a convex function for  $x \geq 0$ .*

Then

$$\mathbf{E}[f(\tilde{X}_n)]^{-1} \sim [f(\mathbf{E}\tilde{X}_n)]^{-1}. \tag{4.5}$$

Denote

$$\tilde{\mu}_{n,s} = \sum_{i=1}^n \mathbf{E}Z_i \mathbf{I}(Z_i \leq \mu_n^s) \quad \text{for some } 0 < s < 1.$$

Consider the following assumption:

- (H<sub>6</sub>)  $\mu_n \sim \tilde{\mu}_{n,s}$  for some  $0 < s < 1$ .

Yang et al. [47] pointed out that condition (H<sub>6</sub>) is weaker than (H<sub>2</sub>) and extended Theorem C for independent random variables to the case of nonnegative random variables under conditions (H<sub>1</sub>) and (H<sub>6</sub>).

Recently, Shen [24] generalized the result of Theorem C to a general case and obtained the following result.

**Theorem E.** *Let  $\{Z_n, n \geq 1\}$  be a sequence of nonnegative random variables with  $\mathbf{E}Z_n < \infty$  for all  $n \geq 1$  and  $0 < s < 1$ . Let  $\{M_n, n \geq 1\}$  and  $\{a_n, n \geq 1\}$  be sequences of positive constants such that  $a_n \geq C$  for all  $n$  sufficiently large, where  $C$  is a positive constant. Denote  $X_n = M_n^{-1} \sum_{k=1}^n Z_k$  and  $\mu_n = \mathbf{E}X_n$  and  $D_n = \eta M_n \mu_n^s / a_n$ , where  $\eta$  is a positive constant. Suppose that the following conditions hold:*

(i) For any  $p > 2$ , there exist positive constants  $\eta$  and  $C$  (depending only on  $p$ ) such that

$$\mathbf{E} \left| \sum_{k=1}^n (Z'_{nk} - \mathbf{E}Z'_{nk}) \right|^p \leq C \left[ \sum_{k=1}^n \mathbf{E}|Z'_{nk} - \mathbf{E}Z'_{nk}|^p + \left( \sum_{k=1}^n \text{Var}(Z'_{nk}) \right)^{p/2} \right],$$

where  $Z'_{nk} = Z_k \mathbf{I}(Z_k \leq D_n) + D_n \mathbf{I}(Z_k > D_n)$ , or  $Z'_{nk} = Z_k \mathbf{I}(Z_k \leq D_n)$ ;

(ii)  $\mu_n \rightarrow \infty$  as  $n \rightarrow \infty$ ;

(iii)

$$\frac{\sum_{k=1}^n \mathbf{E}Z_k \mathbf{I}(Z_k > D_n)}{\sum_{k=1}^n \mathbf{E}Z_k} \rightarrow 0 \quad \text{as } n \rightarrow \infty,$$

where  $\eta > 0$  is the same as that in (i).

Then (4.2) holds for all real constants  $a > 0$  and  $\alpha > 0$ .

Inspired by the literatures above, we will establish the asymptotic approximation of inverse moment as follows.

**Theorem 4.1.** *Let the conditions of Theorem E and (H<sub>3</sub>)–(H<sub>5</sub>) hold. In addition, assume that there exists a positive constant  $\gamma$  such that  $f(x)$  is a nondecreasing function for  $x \geq \gamma$ . Then*

$$\mathbf{E}[f(X_n)]^{-1} \sim [f(\mathbf{E}X_n)]^{-1}. \tag{4.6}$$

*Proof.* Applying Jensen’s inequality to the convex function  $1/f(x)$ , we have

$$\mathbf{E}[f(X_n)]^{-1} \geq [f(\mathbf{E}X_n)]^{-1},$$

which implies that

$$\liminf_{n \rightarrow \infty} f(\mathbf{E}X_n) \mathbf{E}[f(X_n)]^{-1} \geq 1. \tag{4.7}$$

To prove (4.6), we only need to show

$$\limsup_{n \rightarrow \infty} f(\mathbf{E}X_n) \mathbf{E}[f(X_n)]^{-1} \leq 1. \tag{4.8}$$

For any  $0 < \delta < 1$ , let

$$U_n = M_n^{-1} \sum_{k=1}^n [Z_k \mathbf{I}(Z_k \leq D_n) + \eta M_n \mu_n^s / a_n \mathbf{I}(Z_k > D_n)] \doteq M_n^{-1} \sum_{k=1}^n Z'_{nk}$$

and

$$\begin{aligned} \mathbf{E}[f(X_n)]^{-1} &= \mathbf{E}[f(X_n)]^{-1} \mathbf{I}(U_n \geq \mu_n - \delta \mu_n) + \mathbf{E}[f(X_n)]^{-1} \mathbf{I}(U_n < \mu_n - \delta \mu_n) \\ &\doteq Q_1 + Q_2. \end{aligned} \tag{4.9}$$

Note that  $X_n \geq U_n$  and  $f(x)$  is a nondecreasing function for  $x \geq \gamma$ , we have by (H<sub>4</sub>) that

$$\begin{aligned} \limsup_{n \rightarrow \infty} f(\mathbf{E}X_n)Q_1 &\leq \limsup_{n \rightarrow \infty} f(\mu_n)\mathbf{E}[f(U_n)]^{-1}\mathbf{I}(U_n > \mu_n - \delta\mu_n) \\ &\leq \limsup_{n \rightarrow \infty} \left[ \frac{f(\mu_n)}{\mu_n^k} \cdot \frac{\mu_n^k}{(\mu_n - \delta\mu_n)^k} \cdot \frac{(\mu_n - \delta\mu_n)^k}{f(\mu_n - \delta\mu_n)} \right] \\ &= (1 - \delta)^{-k} \rightarrow 1 \quad \text{as } \delta \downarrow 0. \end{aligned} \tag{4.10}$$

In the following, we will prove that

$$\lim_{n \rightarrow \infty} f(\mathbf{E}X_n)Q_2 = 0. \tag{4.11}$$

For  $0 < \delta < 1$  given above, it follows by (iii) in Theorem E that there exists positive integer  $n(\delta) > 0$  such that

$$\sum_{k=1}^n \mathbf{E}Z_k\mathbf{I}(Z_k > D_n) \leq \frac{\delta}{4} \sum_{k=1}^n \mathbf{E}Z_k, \quad n \geq n(\delta), \tag{4.12}$$

which implies that for  $n \geq n(\delta)$ ,

$$\begin{aligned} |\mu_n - \mathbf{E}U_n| &= \left| M_n^{-1} \sum_{k=1}^n \mathbf{E}Z_k\mathbf{I}(Z_k > D_n) - M_n^{-1} \sum_{k=1}^n D_n\mathbf{E}\mathbf{I}(Z_k > D_n) \right| \\ &\leq M_n^{-1} \sum_{k=1}^n \mathbf{E}Z_k\mathbf{I}(Z_k > D_n) + M_n^{-1} \sum_{k=1}^n D_n\mathbf{E}\mathbf{I}(Z_k > D_n) \\ &\leq M_n^{-1} \sum_{k=1}^n \mathbf{E}Z_k\mathbf{I}(Z_k > D_n) + M_n^{-1} \sum_{k=1}^n \mathbf{E}Z_k\mathbf{I}(Z_k > D_n) \\ &= 2M_n^{-1} \sum_{k=1}^n \mathbf{E}Z_k\mathbf{I}(Z_k > D_n) \leq \frac{\delta\mu_n}{2}. \end{aligned} \tag{4.13}$$

By condition (H<sub>3</sub>), (4.13), Markov’s inequality, condition (i) in Theorem E, and  $C_r$ -inequality, we have for any  $p > 2$  and all  $n \geq n(\delta)$  that

$$\begin{aligned} Q_2 &\leq C\mathbf{P}\left(|U_n - \mathbf{E}U_n| > \frac{\delta\mu_n}{2}\right) \leq C\mu_n^{-p}M_n^{-p}\mathbf{E}\left|\sum_{k=1}^n (Z'_{nk} - \mathbf{E}Z'_{nk})\right|^p \\ &\leq C\mu_n^{-p}\left(M_n^{-2}\sum_{k=1}^n \mathbf{E}Z_k^2\mathbf{I}(Z_k \leq D_n) + M_n^{-2}\sum_{k=1}^n D_n^2\mathbf{E}\mathbf{I}(Z_k > D_n)\right)^{p/2} \\ &\quad + C\mu_n^{-p}\left[M_n^{-p}\sum_{k=1}^n \mathbf{E}Z_k^p\mathbf{I}(Z_k \leq D_n) + M_n^{-p}\sum_{k=1}^n D_n^p\mathbf{E}\mathbf{I}(Z_k > D_n)\right] \end{aligned}$$

$$\begin{aligned}
 &\leq C\mu_n^{-p} \left( M_n^{-1} \frac{\mu_n^s}{a_n} \sum_{k=1}^n \mathbf{E}Z_k \mathbf{I}(Z_k \leq D_n) + M_n^{-1} \frac{\mu_n^s}{a_n} \sum_{k=1}^n \mathbf{E}Z_k \mathbf{I}(Z_k > D_n) \right)^{p/2} \\
 &\quad + C\mu_n^{-p} M_n^{-1} \frac{\mu_n^{s(p-1)}}{a_n^{p-1}} \sum_{k=1}^n \mathbf{E}Z_k \mathbf{I}(Z_k \leq D_n) \\
 &\quad + C\mu_n^{-p} M_n^{-1} \frac{\mu_n^{s(p-1)}}{a_n^{p-1}} \sum_{k=1}^n \mathbf{E}Z_k \mathbf{I}(Z_k > D_n) \\
 &= C \left[ \frac{\mu_n^{-(1-s)p/2}}{a_n^{p/2}} + \frac{\mu_n^{-(1-s)(p-1)}}{a_n^{p-1}} \right] \leq C\mu_n^{-(1-s)p/2} + C\mu_n^{-(1-s)(p-1)}. \tag{4.14}
 \end{aligned}$$

Note that  $p > 2$ , and thus  $p - 1 > p/2$ . Taking  $p > \max\{2, 2k/(1 - s)\}$ , we have by (4.14) that  $Q_2 = o(\mu_n^{-k})$ , which together with (H<sub>4</sub>) yield (4.11). Hence, (4.8) follows by (4.9)–(4.11) immediately.

Following similar arguments, we can get (4.6) easily for

$$U_n = M_n^{-1} \sum_{k=1}^n Z_k \mathbf{I}(Z_k \leq D_n) \doteq M_n^{-1} \sum_{k=1}^n Z'_{nk}.$$

This completes the proof of the theorem.

*Remark 4.1.* Since the Rosenthal type inequality (i.e., condition (i) in Theorem E) is satisfied for  $m$ -NOD random variables, the result of Theorem 4.1 holds for nonnegative  $m$ -NOD random variables and other random variables, such as  $\rho$ -mixing random variables,  $\varphi$ -mixing random variables,  $\tilde{\rho}$ -mixing random variables, NA random variables, NSD random variables, NOD random variables, END random variables, AANA random variables,  $\rho^-$ -mixing random variables, and so on.

*Remark 4.2.* Take  $f(x) = (a + x)^\alpha$ ,  $x \geq 0$ ,  $a > 0$ , and  $\alpha > 0$ . It is easy to check that  $f(x)$  is a nondecreasing function for  $x \geq 0$ , and conditions (H<sub>3</sub>)–(H<sub>5</sub>) hold. Hence, Theorem E can be obtained by Theorem 4.1 easily. That is to say, Theorem E is a special case of Theorem 4.1. In addition, if we take  $M_n \equiv 1$  and  $a_n = \eta\sqrt{n}$ , then we can get Theorem D immediately; if we take  $M_n \equiv 1$  and  $a_n = \eta$ , then we can get Theorem 2.3 of [47] immediately; if we take  $a_n = u_n^s$ , then we can get Theorem 2.2 of [9] immediately. Therefore, our result of Theorem 4.1 generalize the corresponding one of [26], [24], [47], and [9].

**Acknowledgments.** The authors are most grateful to the editor and anonymous referee for careful reading of the manuscript and valuable suggestions which helped in significantly improving an earlier version of this paper.

## REFERENCES

1. Vu Thi Ngoc Anh, "A strong limit theorem for sequences of blockwise and pairwise negative quadrant  $m$ -dependent random variables", *Bull. Malays. Math. Sci. Soc.* (2), **36** (2013), 159–164.
2. N. Asadian, V. Fakoor, A. Bozorgnia, "Rosenthal's type inequalities for negatively orthant dependent random variables", *J. Iran. Stat. Soc.*, **5**:1-2 (2006), 69–75.
3. L. E. Baum, M. Katz, "Convergence rates in the law of large numbers", *Trans. Amer. Math. Soc.*, **120**:1 (1965), 108–123.
4. A. Bozorgnia, R. F. Patterson, R. L. Taylor, "Limit theorems for dependent random variables", *World congress of nonlinear analysts '92* (Tampa, FL, 1992), de Gruyter, Berlin, 1996, 1639–1650.
5. M. T. Chao, W. E. Strawderman, "Negative moments of positive random variables", *J. Amer. Statist. Assoc.*, **67**:338 (1972), 429–431.
6. Pingyan Chen, M. O. Cabrera, A. Volodin, " $L_1$ -convergence for weighted sums of some dependent random variables", *Stoch. Anal. Appl.*, **28**:6 (2010), 928–936.
7. T. Fujioka, "Asymptotic approximations of the inverse moment of the non-central chi-squared variable", *J. Japan Statist. Soc.*, **31**:1 (2001), 99–109.
8. N. L. Garcia, J. L. Palacios, "On inverse moments of nonnegative random variables", *Statist. Probab. Lett.*, **53**:3 (2001), 235–239.
9. Shuhe Hu, Xinghui Wang, Wenzhi Yang, Xuejun Wang, "A note on the inverse moment for the non negative random variables", *Comm. Statist. Theory Methods*, **43**:8 (2014), 1750–1757.
10. Tien-Chung Hu, Chen-Yu Chiang, R. L. Taylor, "On complete convergence for arrays of rowwise  $m$ -negatively associated random variables", *Nonlinear Anal.*, **71**:12 (2009), e1075–e1081.
11. Taizhong Hu, "Negatively superadditive dependence of random variables with applications", *Chinese J. Appl. Probab. Statist.*, **16**:2 (2000), 133–144.
12. Taizhong Hu, Jianping Yang, "Further developments on sufficient conditions for negative dependence of random variables", *Statist. Probab. Lett.*, **66**:3 (2004), 369–381.
13. Taizhong Hu, Chaode Xie, Lingyan Ruan, "Dependence structures of multivariate Bernoulli random vectors", *J. Multivariate Anal.*, **94**:1 (2005), 172–195.
14. K. Joag-Dev, F. Proschan, "Negative association of random variables with applications", *Ann. Statist.*, **11**:1 (1983), 286–295.
15. M. Kaluszka, A. Okolewski, "On Fatou-type lemma for monotone moments of weakly convergent random variables", *Statist. Probab. Lett.*, **66**:1 (2004), 45–50.
16. O. Klesov, A. Rosalsky, A. I. Volodin, "On the almost sure growth rate of sums of lower negatively dependent nonnegative random variables", *Statist. Probab. Lett.*, **71**:2 (2005), 193–202.
17. M. Ordóñez Cabrera, A. I. Volodin, "Mean convergence theorems and weak laws of large numbers for weighted sums of random variables under a condition of weighted integrability", *J. Math. Anal. Appl.*, **305**:2 (2005), 644–658.
18. M. Ordóñez Cabrera, A. Rosalsky, A. Volodin, "Some theorems on conditional mean convergence and conditional almost sure convergence for randomly weighted sums of dependent random variables", *Test*, **21**:2 (2012), 369–385.
19. Dehua Qiu, Kuang-Chao Chang, R. Giuliano Antonini, A. Volodin, "On the strong rates of convergence for arrays of rowwise negatively dependent random variables", *Stoch. Anal. Appl.*, **29**:3 (2011), 375–385.

20. Qi-Man Shao, “Maximal inequalities for partial sums of  $\rho$ -mixing sequences”, *Ann. Probab.*, **23**:2 (1995), 948–965.
21. Qi-Man Shao, “A comparison theorem on moment inequalities between negatively associated and independent random variables”, *J. Theoret. Probab.*, **13**:2 (2000), 343–356.
22. Aiting Shen, “Probability inequalities for END sequence and their applications”, *J. Inequal. Appl.*, **2011** (2011), 98, 12 pp.
23. Aiting Shen, “On the strong convergence rate for weighted sums of arrays of rowwise negatively orthant dependent random variables”, *Rev. R. Acad. Cienc. Exactas Fis. Nat. Ser. A Math. RACSAM*, **107**:2 (2013), 257–271.
24. Aiting Shen, “On asymptotic approximation of inverse moments for a class of nonnegative random variables”, *Statistics*, **48**:6 (2014), 1371–1379.
25. Aiting Shen, Ranchao Wu, Yan Chen, Yu Zhou, “Conditional convergence for randomly weighted sums of random variables based on conditional residual  $h$ -integrability”, *J. Inequal. Appl.*, **2013** (2013), 122, 11 pp.
26. Xiaoping Shi, Yuehua Wu, Yu Liu, “A note on asymptotic approximations of inverse moments of nonnegative random variables”, *Statist. Probab. Lett.*, **80**:15–16 (2010), 1260–1264.
27. Xiaoping Shi, N. Reid, Yuehua Wu, “Approximation to the moments of ratios of cumulative sums”, *Canad. J. Statist.*, **42**:2 (2014), 325–336.
28. W. F. Stout, *Almost sure convergence*, Probab. Math. Statist., **24**, Academic Press, New York–London, 1974, x+381 pp.
29. Soo Hak Sung, S. Lisawadi, A. Volodin, “Weak laws of large numbers for arrays under a condition of uniform integrability”, *J. Korean Math. Soc.*, **45**:1 (2008), 289–300.
30. Soo Hak Sung, “Convergence in  $r$ -mean of weighted sums of NQD random variables”, *Appl. Math. Lett.*, **26**:1 (2013), 18–24.
31. R. L. Taylor, R. F. Patterson, A. Bozorgnia, “A strong law of large numbers for arrays of rowwise negatively dependent random variables”, *Stochastic Anal. Appl.*, **20**:3 (2002), 643–656.
32. S. Utev, M. Peligrad, “Maximal inequalities and an invariance principle for a class of weakly dependent random variables”, *J. Theoret. Probab.*, **16**:1 (2003), 101–115.
33. A. Volodin, “On the Kolmogorov exponential inequality for negatively dependent random variables”, *Pakistan J. Statist.*, **18**:2 (2002), 249–253.
34. Jiang Feng Wang, Feng Bin Lu, “Inequalities of maximum of partial sums and weak convergence for a class of weak dependent random variables”, *Acta Math. Sin. (Engl. Ser.)*, **22**:3 (2006), 693–700.
35. Xinghui Wang, Shuhe Hu, “Weak laws of large numbers for arrays of dependent random variables”, *Stochastics*, **86**:5 (2014), 759–775.
36. Xuejun Wang, Shuhe Hu, Wenzhi Yang, Yan Shen, “On complete convergence for weighed sums of  $\varphi$ -mixing random variables”, *J. Inequal. Appl.*, **2010** (2010), 372390, 13 pp.
37. Xuejun Wang, Shuhe Hu, Wenzhi Yang, Nengxiang Ling, “Exponential inequalities and inverse moment for NOD sequence”, *Statist. Probab. Lett.*, **80**:5–6 (2010), 452–461.
38. Xuejun Wang, Shuhe Hu, Wenzhi Yang, “Complete convergence for arrays of rowwise negatively orthant dependent random variables”, *Rev. R. Acad. Cienc. Exactas Fis. Nat. Ser. A Math. RACSAM*, **106**:2 (2012), 235–245.

39. Xuejun Wang, Xin Deng, Lulu Zheng, Shuhe Hu, “Complete convergence for arrays of rowwise negatively superadditive-dependent random variables and its applications”, *Statistics*, **48**:4 (2014), 834–850.
40. Qunying Wu, *Probability limit theory of mixing sequences*, Science Press of China, Beijing, 2006 (in Chinese).
41. Qunying Wu, “A complete convergence theorem for weighted sums of arrays of rowwise negatively dependent random variables”, *J. Inequal. Appl.*, **2012** (2012), 50, 10 pp.
42. Qunying Wu, Yuanying Jiang, “The strong consistency of  $M$  estimator in a linear model for negatively dependent random samples”, *Comm. Statist. Theory Methods*, **40**:3 (2011), 467–491.
43. Tiejie-Jian Wu, Xiaoping Shi, Baiqi Miao, “Asymptotic approximation of inverse moments of nonnegative random variables”, *Statist. Probab. Lett.*, **79**:11 (2009), 1366–1371.
44. Yongfeng Wu, Mei Guan, “Mean convergence theorems and weak laws of large numbers for weighted sums of dependent random variables”, *J. Math. Anal. Appl.*, **377**:2 (2011), 613–623.
45. Yongfeng Wu, M. Ordóñez Cabrera, A. Volodin, “On limiting behavior for arrays of rowwise negatively orthant dependent random variables”, *J. Korean Statist. Soc.*, **42** (2013), 61–70.
46. Yongfeng Wu, A. Rosalsky, “Strong convergence for  $m$ -pairwise negatively quadrant dependent random variables”, *Glas. Mat. Ser. III*, **50(70)**:1 (2015), 245–259.
47. Wenzhi Yang, Shuhe Hu, Xuejun Wang, “On the asymptotic approximation of inverse moment for non negative random variables”, *Comm. Statist. Theory Methods*, **46**:16 (2017), 7787–7797.
48. Demei Yuan, Bao Tao, “Mean convergence theorems for weighted sums of arrays of residually  $h$ -integrable random variables concerning the weights under dependence assumptions”, *Acta Appl. Math.*, **103**:3 (2008), 221–234.
49. DeMei Yuan, Jun An, “Rosenthal type inequalities for asymptotically almost negatively associated random variables and applications”, *Sci. China Ser. A*, **52**:9 (2009), 1887–1904.
50. H. Zarei, H. Jabbari, “Complete convergence of weighted sums under negative dependence”, *Statist. Papers*, **52**:2 (2011), 413–418.
51. L.-X. Zhang, “A functional central limit theorem for asymptotically negatively dependent random fields”, *Acta Math. Hungar.*, **86**:3 (2000), 237–259.

Поступила в редакцию  
31.III.2015