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# STRONG CONVERGENCE PROPERTIES FOR ARRAYS OF ROWWISE NEGATIVELY ORTHANT DEPENDENT RANDOM VARIABLES

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ABSTRACT. Let  $\{X_{ni}, i \geq 1, n \geq 1\}$  be an array of rowwise negatively orthant dependent random variables which is stochastically dominated by a random variable X. Wang et al. [15. Complete convergence for arrays of rowwise negatively orthant dependent random variables, RACSAM, **106** (2012), 235–245] studied the complete convergence for arrays of rowwise negatively orthant dependent random variables under the condition that X has an exponential moment, which seems too strong. We will further study the complete convergence for arrays of rowwise negatively orthant dependent random variables under the condition that X has a moment, which is weaker than exponential moment. Our results improve the corresponding ones of Wang et al. [15].

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## 1. Introduction

Let  $\{X, X_n, n \ge 1\}$  be a sequence of identically distributed random variables and  $\{a_{ni}, 1 \le i \le n, n \ge 1\}$  an array of constants. The strong convergence results for weighted sums  $\sum_{i=1}^{n} a_{ni}X_i$  have been studied by many authors, see for example, Bai and Cheng [2], Cai [3], Chen and Gan [6], Cuzick [7], Sung [13], Wang et al. [15–18], Wu [19–21], Zhou et al. [27], Xu and Tang [25, 26], Wu et al. [22], Tang [14] and so forth. Many useful linear statistics are these weighted sums. Examples include least squares estimators, nonparametric regression function estimators and jackknife estimates among others. Bai and Cheng [2] proved the strong law of large numbers for weighted sums

$$\frac{1}{b_n}\sum_{i=1}^n a_{ni}X_i \to 0, \quad a.s.$$

when  $\{X, X_n, n \ge 1\}$  is a sequence of independent and identically distributed random variables with EX = 0 and  $E \exp(h|X|^{\gamma}) < \infty$  for some h > 0 and  $\gamma > 0$ , and  $\{a_{ni}, 1 \le i \le n, n \ge 1\}$  is an

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array of constants satisfying

$$A_{\alpha} \doteq \limsup_{n \to \infty} A_{\alpha,n} < \infty, \qquad A_{\alpha,n}^{\alpha} \doteq \frac{1}{n} \sum_{i=1}^{n} |a_{ni}|^{\alpha}$$

for some  $1 < \alpha < 2$ , where  $b_n = n^{1/\alpha} (\log n)^{1/\gamma + \gamma(\alpha - 1)/\alpha(1+\gamma)}$ .

The concept of complete convergence was introduced by Hsu and Robbins [10] as follows: a sequence of random variables  $\{U_n, n \geq 1\}$  is said to converge completely to a constant C if  $\sum_{n=1}^{\infty} P(|U_n - C| > \varepsilon) < \infty$  for all  $\varepsilon > 0$ . In view of the Borel-Cantelli lemma, this implies that  $U_n \to C$  almost surely (a.s.). The converse is true if the  $\{U_n, n \geq 1\}$  are independent. Hsu and Robbins [10] proved that the sequence of arithmetic means of independent and identically distributed (i.i.d.) random variables converges completely to the expected value if the variance of the summands is finite. Erdös [8] proved the converse. The result of Hsu-Robbins-Erdös is a fundamental theorem in probability theory and has been extended in several directions by many authors. One of the most important generalizations is the Baum-Katz-Spitzer type result. For more details about the Baum-Katz-Spitzer type results, one can refer to Spitzer [12], Baum and Katz [4] and Gut [9], and so forth.

A finite collection of random variables  $X_1, X_2, \ldots, X_n$  is said to be negatively associated (NA) if for every pair of disjoint subsets  $A_1, A_2$  of  $\{1, 2, \ldots, n\}$ ,

$$Cov\{f(X_i : i \in A_1), g(X_j : j \in A_2)\} \le 0,$$

whenever f and g are coordinatewise nondecreasing such that this covariance exists. An infinite sequence  $\{X_n, n \ge 1\}$  is NA if every finite subcollection is NA.

An array of random variables  $\{X_{ni}, i \ge 1, n \ge 1\}$  is called rowwise NA random variables if for every  $n \ge 1$ ,  $\{X_{ni}, i \ge 1\}$  is a sequence of NA random variables.

A finite collection of random variables  $X_1, X_2, \ldots, X_n$  is said to be negatively orthant dependent (NOD) if

$$P(X_1 > x_1, X_2 > x_2, \dots, X_n > x_n) \le \prod_{i=1}^n P(X_i > x_i)$$

and

$$P(X_1 \le x_1, X_2 \le x_2, \dots, X_n \le x_n) \le \prod_{i=1}^n P(X_i \le x_i)$$

hold for all  $x_1, x_2, \ldots, x_n \in \mathbb{R}$ . An infinite sequence  $\{X_n, n \ge 1\}$  is said to be NOD if every finite subcollection is NOD.

An array of random variables  $\{X_{ni}, i \ge 1, n \ge 1\}$  is called rowwise NOD if for every  $n \ge 1$ ,  $\{X_{ni}, i \ge 1\}$  is a sequence of NOD random variables.

The concepts of NA and NOD sequences were introduced by Joag-Dev and Proschan [11]. Obviously, independent random variables are NOD. Joag-Dev and Proschan [11] pointed out that NOD is weaker than NA.

Recently, Cai [3] obtained the following complete convergence result for weighted sums of identically distributed NA random variables.

**THEOREM 1.1.** Let  $\{X, X_n, n \ge 1\}$  be a sequence of NA random variables with identical distribution, and  $\{a_{ni}, 1 \le i \le n, n \ge 1\}$  be a triangular array of constants satisfying  $\sum_{i=1}^{n} |a_{ni}|^{\alpha} = O(n)$ for  $0 < \alpha \le 2$ . Let  $b_n = n^{1/\alpha} \log^{1/\gamma} n$  for some  $\gamma > 0$ . Furthermore, assume that EX = 0 when  $1 < \alpha \leq 2$ . If  $E \exp(h|X|^{\gamma}) < \infty$  for some h > 0, then for any  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} \frac{1}{n} P\left( \max_{1 \le j \le n} \left| \sum_{i=1}^{j} a_{ni} X_i \right| > \varepsilon b_n \right) < \infty.$$

Wang et al. [15] extended the result of Cai [3] for sequences of NA random variables to the case of arrays of rowwise NOD random variables and obtained the following result.

**THEOREM 1.2.** Let  $\{X_{ni} : i \geq 1, n \geq 1\}$  be an array of rowwise NOD random variables which is stochastically dominated by a random variable X and  $\{a_{ni} : i \geq 1, n \geq 1\}$  be an array of real numbers. Assume that there exist some  $\delta$  with  $0 < \delta < 1$  and some  $\alpha$  with  $0 < \alpha \leq 2$  such that  $\sum_{i=1}^{n} |a_{ni}|^{\alpha} = O(n^{\delta})$  and assume further that  $EX_{ni} = 0$  when  $1 < \alpha \leq 2$ . If for some h > 0 and  $\gamma > 0$  such that  $E \exp(h|X|^{\gamma}) < \infty$ , then for any  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} n^{p\alpha-2} P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} a_{ni} X_{ni} \right| > \varepsilon b_n \right) < \infty,$$

where  $p \geq 1/\alpha$  and  $b_n \doteq n^{1/\alpha} \log^{1/\gamma} n$ .

All the results above are based on the condition that  $E \exp(h|X|^{\gamma}) < \infty$  for some h > 0 and  $\gamma > 0$  (or for all h > 0 and some  $\gamma > 0$ ). The exponential moment seems too strong. The question is that whether the exponential moment can be replaced by a moment, i.e., there exists a constant  $\beta > 0$  such that  $E|X|^{\beta} < \infty$ . Our answer is positive.

Our goal in this paper is to further study the complete convergence for arrays of rowwise NOD random variables under the condition that X has a moment, which is weaker than exponential moment. The results presented in this paper are inspired by Wang et al. [15]. The techniques used in the paper are the truncated method and the Rosenthal type inequality for NOD random variables.

**DEFINITION 1.1.** An array of random variables  $\{X_{ni}, i \ge 1, n \ge 1\}$  is said to be stochastically dominated by a random variable X if there exists a positive constant C such that

$$P(|X_{ni}| > x) \le CP(|X| > x)$$

for all  $x \ge 0$ ,  $i \ge 1$  and  $n \ge 1$ .

The following lemmas are useful for the proofs of the main results. The first one is a basic property for NOD random variables, which was given by Bozorgnia et al. [5].

**LEMMA 1.1.** ([5]) Let random variables  $X_1, X_2, \ldots, X_n$  be NOD,  $f_1, f_2, \ldots, f_n$  be all nondecreasing (or all nonincreasing) functions, then random variables  $f_1(X_1), f_2(X_2), \ldots, f_n(X_n)$  are NOD.

The next one is the Rosenthal type inequality for NOD random variables. For the proofs, one can refer to Asadian et al. [1] and Wu [24].

**LEMMA 1.2.** ([1,24]) Let  $p \ge 2$  and  $\{X_n, n \ge 1\}$  be a sequence of NOD random variables with  $EX_n = 0$  and  $E|X_n|^p < \infty$  for every  $n \ge 1$ . Then there exists a positive constant C depending only on p such that for every  $n \ge 1$ ,

$$E\left|\sum_{i=1}^{n} X_{i}\right|^{p} \leq C\left\{\sum_{i=1}^{n} E|X_{i}|^{p} + \left(\sum_{i=1}^{n} EX_{i}^{2}\right)^{p/2}\right\},\$$
$$E\left(\max_{1\leq j\leq n}\left|\sum_{i=1}^{j} X_{i}\right|^{p}\right) \leq C\log^{p} 2n\left\{\sum_{i=1}^{n} E|X_{i}|^{p} + \left(\sum_{i=1}^{n} EX_{i}^{2}\right)^{p/2}\right\}.$$

The last one is a basic property for stochastic domination. For the proof, one can refer to Wu [24] or Wang et al. [23].

**LEMMA 1.3.** Let  $\{X_n, n \ge 1\}$  be a sequence of random variables which is stochastically dominated by a random variable X. Then for any  $\alpha > 0$  and b > 0,

$$E|X_n|^{\alpha}I(|X_n| \le b) \le C_1[E|X|^{\alpha}I(|X| \le b) + b^{\alpha}P(|X| > b)]$$

$$E|X_n|^{\alpha}I(|X_n| > b) \le C_2 E|X|^{\alpha}I(|X| > b)$$

where  $C_1$  and  $C_2$  are positive constants.

Throughout the paper, let I(A) be the indicator function of the set A. C denotes a positive constant which may be different in various places and  $a_n = O(b_n)$  stands for  $a_n \leq Cb_n$ .

### 2. Main results and their proofs

Our main results are as follows.

**THEOREM 2.1.** Let  $\{X_{ni}, i \geq 1, n \geq 1\}$  be an array of rowwise NOD random variables which is stochastically dominated by a random variable X and  $\{a_{ni}, i \geq 1, n \geq 1\}$  be an array of real numbers. Assume that the following two conditions are satisfied:

(i) There exist some  $\delta$  with  $0 < \delta < 1$  and some  $\alpha$  with  $0 < \alpha \leq 2$  such that  $\sum_{i=1}^{n} |a_{ni}|^{\alpha} = O(n^{\delta})$ and assume further that  $EX_{ni} = 0$  when  $1 < \alpha \leq 2$ ;

(ii)  $p \ge 1/\alpha$ . For some  $\beta > \max\{p\alpha^2, \alpha + \frac{\alpha(p\alpha-1)}{1-\delta}, \alpha+2, \alpha(p\alpha-1)+2\delta\}, E|X|^{\beta} < \infty$ . Then for any  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} n^{p\alpha-2} P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} a_{ni} X_{ni} \right| > \varepsilon b_n \right) < \infty,$$

$$(2.1)$$

where  $b_n \doteq n^{1/\alpha} \log^{1/\gamma} n$  for some  $\gamma > 0$ .

Proof. For fixed  $n \ge 1$ , define

$$X_{i}^{(n)} = -b_{n}I(X_{ni} < -b_{n}) + X_{ni}I(|X_{ni}| \le b_{n}) + b_{n}I(X_{ni} > b_{n}), \quad i \ge 1$$
$$T_{j}^{(n)} = \sum_{i=1}^{j} a_{ni} \left( X_{i}^{(n)} - EX_{i}^{(n)} \right), \quad j = 1, 2, \dots, n.$$

It is easy to check that for any  $\varepsilon > 0$ ,

$$\left(\max_{1\leq j\leq n}\left|\sum_{i=1}^{j}a_{ni}X_{ni}\right| > \varepsilon b_{n}\right) \subset \left(\max_{1\leq i\leq n}\left|X_{ni}\right| > b_{n}\right) \cup \left(\max_{1\leq j\leq n}\left|\sum_{i=1}^{j}a_{ni}X_{i}^{(n)}\right| > \varepsilon b_{n}\right), \quad (2.2)$$

which implies that

$$P\left(\max_{1\leq j\leq n} \left|\sum_{i=1}^{j} a_{ni} X_{ni}\right| > \varepsilon b_{n}\right)$$

$$\leq P\left(\max_{1\leq i\leq n} |X_{ni}| > b_{n}\right) + P\left(\max_{1\leq j\leq n} \left|\sum_{i=1}^{j} a_{ni} X_{i}^{(n)}\right| > \varepsilon b_{n}\right)$$

$$\leq \sum_{i=1}^{n} P\left(|X_{ni}| > b_{n}\right) + P\left(\max_{1\leq j\leq n} \left|T_{j}^{(n)}\right| > \varepsilon b_{n} - \max_{1\leq j\leq n} \left|\sum_{i=1}^{j} a_{ni} E X_{i}^{(n)}\right|\right).$$

$$(2.3)$$

Firstly, we will show that

$$b_n^{-1} \max_{1 \le j \le n} \left| \sum_{i=1}^j a_{ni} E X_i^{(n)} \right| \to 0, \quad \text{as} \quad n \to \infty.$$

$$(2.4)$$

By  $\sum_{i=1}^{n} |a_{ni}|^{\alpha} = O(n^{\delta})$  and Hölder's inequality, we have for  $1 \le k < \alpha$  that

$$\sum_{i=1}^{n} |a_{ni}|^k \le \left(\sum_{i=1}^{n} \left(|a_{ni}|^k\right)^{\frac{\alpha}{k}}\right)^{\frac{k}{\alpha}} \left(\sum_{i=1}^{n} 1\right)^{\frac{\alpha-k}{\alpha}} \le Cn.$$
(2.5)

Hence, when  $1 < \alpha \leq 2$ , we have by  $EX_{ni} = 0$ , Lemma 1.3, (2.5) (Taking k = 1), Markov's inequality and condition (ii) that

$$\begin{split} b_{n}^{-1} \max_{1 \leq j \leq n} \left| \sum_{i=1}^{j} a_{ni} E X_{i}^{(n)} \right| \\ &\leq \sum_{i=1}^{n} |a_{ni}| P(|X_{ni}| > b_{n}) + b_{n}^{-1} \max_{1 \leq j \leq n} \left| \sum_{i=1}^{j} a_{ni} E X_{ni} I(|X_{ni}| > b_{n}) \right| \\ &\leq C \sum_{i=1}^{n} |a_{ni}| P(|X| > b_{n}) + b_{n}^{-1} \sum_{i=1}^{n} |a_{ni}| E |X_{ni}| I(|X_{ni}| > b_{n}) \\ &\leq C n \frac{E|X|^{\beta}}{b_{n}^{\beta}} + C b_{n}^{-1} \sum_{i=1}^{n} |a_{ni}| E|X| I(|X| > b_{n}) \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n E|X| I(|X| > b_{n}) \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} E|X| I(b_{k} < |X| \leq b_{k+1}) \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} P(|X| > b_{k}) \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} n \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\alpha}}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\gamma}}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\gamma}}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} + C b_{n}^{-1} \sum_{k=n}^{\infty} b_{k+1} \frac{E|X|^{\beta}}{b_{k}^{\beta/\gamma}}} \\ &\leq \frac{C n}{n^{\beta/\alpha} \log^{\beta$$

Elementary Jensen's inequality implies that for any 0 < s < t,

$$\left(\sum_{i=1}^{n} |a_{ni}|^t\right)^{1/t} \le \left(\sum_{i=1}^{n} |a_{ni}|^s\right)^{1/s}.$$
(2.7)

Therefore, when  $0 < \alpha \leq 1$ , we have by Lemma 1.3, (2.7), Markov's inequality and condition (ii) that

$$\begin{split} b_{n}^{-1} \max_{1 \le j \le n} \left| \sum_{i=1}^{j} a_{ni} E X_{i}^{(n)} \right| \\ &\leq \sum_{i=1}^{n} |a_{ni}| P(|X_{ni}| > b_{n}) + b_{n}^{-1} \sum_{i=1}^{n} |a_{ni}| E|X_{ni}| I(|X_{ni}| \le b_{n}) \\ &\leq C \sum_{i=1}^{n} |a_{ni}| P(|X| > b_{n}) + C b_{n}^{-1} \sum_{i=1}^{n} |a_{ni}| \left( E|X|I(|X| \le b_{n}) + b_{n}P(|X| > b_{n}) \right) \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} E|X|I(|X| \le b_{n}) + C n^{\delta/\alpha} P(|X| > b_{n}) \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} E|X|I(b_{k-1} < |X| \le b_{k}) + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{b_{n}^{\beta}} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} b_{k} P(|X| > b_{k-1}) + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} b_{k} \frac{E|X|^{\beta}}{b_{k-1}^{\beta-1}} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} \frac{k^{1/\alpha} \log^{1/\gamma} k}{(k-1)^{\beta/\alpha} \log^{\beta/\gamma} (k-1)} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha - \beta/\alpha} + \frac{C n^{\delta/\alpha} E|X|^{\beta}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha} \sum_{k=2}^{n} k^{1/\alpha} \sum_{k=2}^{n} k^{1/\alpha} \sum_{k=2}^{n} k^{1/\alpha} \\ &\leq C b_{n}^{-1} n^{\delta/\alpha} \sum_{k=2}^{n} k^{1/\alpha} \sum_{k=2}^{n} k^{1/\alpha} \sum_{k$$

By (2.6) and (2.8), we can get (2.4) immediately. Hence, for n large enough,

$$P\Big(\max_{1\leq j\leq n}\Big|\sum_{i=1}^{j}a_{ni}X_{ni}\Big|>\varepsilon b_n\Big)\leq \sum_{i=1}^{n}P\big(|X_{ni}|>b_n\big)+P\Big(\max_{1\leq j\leq n}\big|T_j^{(n)}\big|>\frac{\varepsilon}{2}b_n\Big).$$

To prove (2.1), we only need to show that

$$I \doteq \sum_{n=1}^{\infty} n^{p\alpha-2} \sum_{i=1}^{n} P\Big(|X_{ni}| > b_n\Big) < \infty$$

$$(2.9)$$

and

$$J \doteq \sum_{n=1}^{\infty} n^{p\alpha-2} P\left(\max_{1 \le j \le n} \left| T_j^{(n)} \right| > \frac{\varepsilon}{2} b_n \right) < \infty.$$

$$(2.10)$$

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By Definition 1.1, Markov's inequality and condition (ii), we can see that

$$I \doteq \sum_{n=1}^{\infty} n^{p\alpha-2} \sum_{i=1}^{n} P\left(|X_{ni}| > b_n\right)$$
  
$$\leq C \sum_{n=1}^{\infty} n^{p\alpha-2} \sum_{i=1}^{n} P\left(|X| > b_n\right) \leq C \sum_{n=2}^{\infty} n^{p\alpha-1} \frac{E|X|^{\beta}}{b_n^{\beta}}$$
  
$$\leq C \sum_{n=2}^{\infty} \frac{n^{p\alpha-1}}{n^{\beta/\alpha} \log^{\beta/\gamma} n} < \infty, \quad (\text{since } \beta > p\alpha^2).$$
  
(2.11)

For fixed  $n \ge 1$ , it is easily seen that  $\{X_i^{(n)}, 1 \le i \le n\}$  are still NOD by Lemma 1.1. For q > 2, it follows from Lemma 1.2,  $C_r$  inequality and Jensen's inequality that

$$J \doteq \sum_{n=1}^{\infty} n^{p\alpha-2} P\left(\max_{1 \le j \le n} \left| T_j^{(n)} \right| > \frac{\varepsilon}{2} b_n \right) \le C \sum_{n=2}^{\infty} n^{p\alpha-2} b_n^{-q} E\left(\max_{1 \le j \le n} \left| T_j^{(n)} \right|^q\right)$$
$$\le C \sum_{n=2}^{\infty} n^{p\alpha-2} b_n^{-q} (\log n)^q \left[ \sum_{i=1}^n |a_{ni}|^q E \left| X_i^{(n)} \right|^q + \left( \sum_{i=1}^n |a_{ni}|^2 E \left| X_i^{(n)} \right|^2 \right)^{q/2} \right]$$
$$\doteq J_1 + J_2.$$
(2.12)

Taking a suitable constant q such that  $\max\{2, \alpha(p\alpha - 1)/(1 - \delta)\} < q < \min\{\beta - \alpha, \frac{\beta - p\alpha^2 + \alpha}{\delta}\}$ , which implies that

$$\beta > \alpha + q, \quad \frac{\beta}{\alpha} - \frac{q}{\alpha} > 1, \quad \beta > p\alpha^2 - \alpha + q\delta, \quad \frac{\beta}{\alpha} - p\alpha + 2 - q\frac{\delta}{\alpha} > 1$$

and

$$p\alpha - 2 + q\frac{\delta}{\alpha} - \frac{q}{\alpha} < -1, \qquad q > \alpha.$$

It follows from  $C_r$  inequality, Lemma 1.3, (2.7), Markov's inequality and condition (ii) that

$$J_{1} \doteq C \sum_{n=2}^{\infty} n^{p\alpha-2} b_{n}^{-q} (\log n)^{q} \sum_{i=1}^{n} |a_{ni}|^{q} E \left| X_{i}^{(n)} \right|^{q}$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2} b_{n}^{-q} (\log n)^{q} \sum_{i=1}^{n} |a_{ni}|^{q} \left[ E \left| X_{ni} \right|^{q} I(|X_{ni}| \le b_{n}) + b_{n}^{q} P(|X_{ni}| > b_{n}) \right]$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2} b_{n}^{-q} (\log n)^{q} \sum_{i=1}^{n} |a_{ni}|^{q} \left[ E |X|^{q} I(|X| \le b_{n}) + b_{n}^{q} P(|X| > b_{n}) \right]$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} b_{n}^{-q} (\log n)^{q} E |X|^{q} I(|X| \le b_{n})$$

$$+ C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} (\log n)^{q} P(|X| > b_{n})$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} b_{n}^{-q} (\log n)^{q} \sum_{k=2}^{n} E |X|^{q} I(b_{k-1} < |X| \le b_{k})$$

$$+ C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} (\log n)^{q} \frac{E |X|^{\beta}}{b_{n}^{\beta}}$$

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$$\leq C \sum_{k=2}^{\infty} \sum_{n=k}^{\infty} n^{p\alpha-2+q\delta/\alpha-q/\alpha} (\log n)^{q-q/\gamma} b_k^q P(|X| > b_{k-1}) + C \sum_{n=2}^{\infty} \frac{n^{p\alpha-2+q\delta/\alpha} \log^q n}{n^{\beta/\alpha} \log^{\beta/\gamma} n} \leq C \sum_{k=3}^{\infty} b_k^q \frac{E|X|^{\beta}}{b_{k-1}^{\beta}} + C \sum_{n=2}^{\infty} \frac{1}{n^{\beta/\alpha-p\alpha+2-q\delta/\alpha} \log^{\beta/\gamma-q} n} \leq C \sum_{k=3}^{\infty} \frac{k^{q/\alpha} \log^{q/\gamma} k}{(k-1)^{\beta/\alpha} \log^{\beta/\gamma} (k-1)} + C \sum_{n=2}^{\infty} \frac{1}{n^{\beta/\alpha-p\alpha+2-q\delta/\alpha} \log^{\beta/\gamma-q} n} \leq C \sum_{k=3}^{\infty} \frac{1}{k^{\beta/\alpha-q/\alpha}} + C \sum_{n=2}^{\infty} \frac{1}{n^{\beta/\alpha-p\alpha+2-q\delta/\alpha} \log^{\beta/\gamma-q} n} < \infty.$$

$$(2.13)$$

By  $C_r$  inequality, Lemma 1.3, (2.7) and Jensen's inequality, we can get that

$$J_{2} \doteq C \sum_{n=2}^{\infty} n^{p\alpha-2} b_{n}^{-q} (\log n)^{q} \left( \sum_{i=1}^{n} |a_{ni}|^{2} E \left| X_{i}^{(n)} \right|^{2} \right)^{q/2}$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2} b_{n}^{-q} (\log n)^{q} \left( \sum_{i=1}^{n} |a_{ni}|^{2} \left[ E |X_{ni}|^{2} I(|X_{ni}| \le b_{n}) + b_{n}^{2} P(|X_{ni}| > b_{n}) \right] \right)^{q/2}$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2} b_{n}^{-q} (\log n)^{q} \left( \sum_{i=1}^{n} |a_{ni}|^{2} \left[ E X^{2} I(|X| \le b_{n}) + b_{n}^{2} P(|X| > b_{n}) \right] \right)^{q/2}$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} b_{n}^{-q} (\log n)^{q} \left[ E X^{2} I(|X| \le b_{n}) + b_{n}^{2} P(|X| > b_{n}) \right]^{q/2}$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} b_{n}^{-q} (\log n)^{q} \left[ E X^{2} I(|X| \le b_{n}) + b_{n}^{2} P(|X| > b_{n}) \right]^{q/2}$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} b_{n}^{-q} (\log n)^{q} \left[ P(|X| > b_{n}) \right]^{q/2}$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} (\log n)^{q} \left[ P(|X| > b_{n}) \right]^{q/2}$$

$$\leq C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} b_{n}^{-q} (\log n)^{q} E |X|^{q} I(|X| \le b_{n})$$

$$+ C \sum_{n=2}^{\infty} n^{p\alpha-2+q\delta/\alpha} (\log n)^{q} P(|X| > b_{n}) < \infty. \quad (by (2.13))$$

Therefore, the desired result (2.1) follows from (2.11)–(2.14) immediately. This completes the proof of the theorem.  $\hfill \Box$ 

Similar to the proof of Theorem 2.1 above and Theorems 2.3–2.6 of Wang et al. [15], we can get the following results.

**THEOREM 2.2.** Let  $\{X_n, n \ge 1\}$  be a sequence of NOD random variables which is stochastically dominated by a random variable X and  $\{a_n, n \ge 1\}$  be a sequence of real numbers. Assume that there exist some  $\delta$  with  $0 < \delta < 1$  and some  $\alpha$  with  $0 < \alpha \le 2$  such that  $\sum_{i=1}^{n} |a_i|^{\alpha} = O(n^{\delta})$  and assume further that  $EX_n = 0$  when  $1 < \alpha \le 2$ . If condition (ii) of Theorem 2.1 holds, then for any  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} n^{p\alpha-2} P\left(\max_{1 \le j \le n} \left| S_j \right| > \varepsilon b_n\right) < \infty$$

and

$$\lim_{n \to \infty} \frac{|S_n|}{b_n} = 0 \quad a.s.,$$

where  $b_n \doteq n^{1/\alpha} \log^{1/\gamma} n$  for some  $\gamma > 0$  and  $S_n = \sum_{i=1}^n a_i X_i$  for  $n \ge 1$ .

**THEOREM 2.3.** Let  $\{X_{ni} : i \ge 1, n \ge 1\}$  be an array of rowwise NOD random variables which is stochastically dominated by a random variable X and  $\{a_{ni} : i \ge 1, n \ge 1\}$  be an array of real numbers. Assume that there exists some  $\alpha$  with  $0 < \alpha \le 2$  such that  $\sum_{i=1}^{n} |a_{ni}|^{\alpha} = O(n)$  and assume further that  $EX_{ni} = 0$  when  $1 < \alpha \le 2$ . If there exists some  $\beta > \alpha + 2$  such that  $E|X|^{\beta} < \infty$ , then for any  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} n^{-1} P\left( \max_{1 \le j \le n} \left| \sum_{i=1}^{j} a_{ni} X_{ni} \right| > \varepsilon b_n \right) < \infty,$$

where  $b_n \doteq n^{1/\alpha} \log^{1/\gamma} n$  for some  $\gamma > 0$ .

**THEOREM 2.4.** Let  $\{X_n, n \ge 1\}$  be a sequence of NOD random variables which is stochastically dominated by a random variable X and  $\{a_{ni}, i \ge 1, n \ge 1\}$  be an array of real numbers. Assume that there exists some  $\alpha$  with  $0 < \alpha \le 2$  such that  $\sum_{i=1}^{n} |a_{ni}|^{\alpha} = O(n)$  and assume further that  $EX_n = 0$  when  $1 < \alpha \le 2$ . If there exists some  $\beta > \alpha + 2$  such that  $E|X|^{\beta} < \infty$ , then for any  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} n^{-1} P\bigg( \max_{1 \le j \le n} \bigg| \sum_{i=1}^{j} a_{ni} X_i \bigg| > \varepsilon b_n \bigg) < \infty,$$

where  $b_n \doteq n^{1/\alpha} \log^{1/\gamma} n$  for some  $\gamma > 0$ .

**THEOREM 2.5.** Let  $\{X_n, n \ge 1\}$  be a sequence of NOD random variables which is stochastically dominated by a random variable X and  $\{a_n, n \ge 1\}$  be a sequence of real numbers. Assume that there exists some  $\alpha$  with  $0 < \alpha \le 2$  such that  $\sum_{i=1}^{n} |a_i|^{\alpha} = O(n)$  and assume further that  $EX_n = 0$  when  $1 < \alpha \le 2$ . If there exists some  $\beta > \alpha + 2$  such that  $E|X|^{\beta} < \infty$ , then for any  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} n^{-1} P\left( \max_{1 \le j \le n} \left| S_j \right| > \varepsilon b_n \right) < \infty$$

and

$$\lim_{n \to \infty} \frac{|S_n|}{b_n} = 0 \quad a.s.$$

where  $b_n \doteq n^{1/\alpha} \log^{1/\gamma} n$  for some  $\gamma > 0$  and  $S_n = \sum_{i=1}^n a_i X_i$  for  $n \ge 1$ .

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#### REFERENCES

- ASADIAN, N.—FAKOOR, V.—BOZORGNIA, A.: Rosenthal's type inequalities for negatively orthant dependent random variables, Iran. Stat. Soc. (JIRSS) 5 (2006), 66–75.
- [2] BAI, Z. D.—CHENG, P. E.: Marcinkiewicz strong laws for linear statistics, Statist. Prob. Lett. 46 (2000), 105–112.
- [3] CAI, G. H.: Strong laws for weighted sums of NA random variables, Metrika 68 (2008), 323-331.

- [4] BAUM, L. E.—KATZ, M.: Convergence rates in the law of large numbers, Trans. Amer. Math. Soc. 120 (1965), 108–123.
- BOZORGNIA, A.—PATTERSON, R. F.—TAYLOR R. L.: Limit theorems for dependent random variables, World Congress Nonlinear Analysts, 92 (1996), 1639–1650.
- [6] CHEN, P. Y.—GAN, S. X.: Limiting behavior of weighted sums of i.i.d. random variables, Statist. Probab. Lett. 77 (2007), 1589–1599.
- [7] CUZICK, J.: A strong law for weighted sums of i.i.d. random variables, J. Theoret. Probab. 8 (1995), 625-641.
- [8] ERDÖS, P.: On a theorem of Hsu and Robbins, Ann. Math. Statist. 20 (1949), 286-291.
- [9] GUT, A.: Complete convergence for arrays, Period. Math. Hungar. 25 (1992), 51-75.
- [10] HSU, P. L.—ROBBINS, H.: Complete convergence and the law of large numbers, Proc. Natl. Acad. Sci. USA 33 (1947), 25–31.
- [11] JOAG-DEV, K.—PROSCHAN, F.: Negative association of random variables with applications, Ann. Statist. 11 (1983), 286–295.
- [12] SPITZER, F. L.: A combinatorial lemma and its application to probability theory, Trans. Amer. Math. Soc. 82 (1956), 323–339.
- [13] SUNG, S. H.: On the strong convergence for weighted sums of random variables, Statist. Papers 52 (2011), 447–454.
- [14] TANG, X. F.: Some strong laws of large numbers for weighted sums of asymptotically almost negatively associated random variables, J. Inequal. Appl. 2013 (2013), Article ID 4.
- [15] WANG, X. J.—HU, S. H.—YANG, W. Z.: Complete convergence for arrays of rowwise negatively orthant dependent random variables, Rev. R. Acad. Cienc. Exactas Fís. Nat. Ser. A Math. RACSAM 106 (2012), 235–245.
- [16] WANG, X. J.—LI, X. Q.—YANG, W. Z.—HU, S. H.: On complete convergence for arrays of rowwise weakly dependent random variables, Appl. Math. Lett. 25 (2012), 1916–1920.
- [17] WANG, X. J.—DENG, X.—ZHENG, L. L.—HU, S. H.: Complete convergence for arrays of rowwise negatively superadditive-dependent random variables and its applications, Statistics 48 (2014), 834–850.
- [18] WANG, X. J.—XU, C.—HU, T. C.—VOLODIN, A.—HU, C.: On complete convergence for widely orthantdependent random variables and its applications in nonparametric regression models, TEST 23 (2014), 607–629.
- [19] WU, Q.Y.: A strong limit theorem for weighted sums of sequences of negatively dependent random variables, J. Inequal. Appl. 2010 (2010), Article ID 383805.
- [20] WU, Q. Y.: Sufficient and necessary conditions of complete convergence for weighted sums of PNQD random variables, J. Appl. Math. 2012 (2012), Article ID 104390.
- [21] WU, Q. Y.: A complete convergence theorem for weighted sums of arrays of rowwise negatively dependent random variables, J. Inequal. Appl. 2012 (2012), Article ID 50.
- [22] WU, Y. F.-WANG, C. H.-VOLODIN, A.: Limiting behavior for arrays of rowwise ρ<sup>\*</sup>-mixing random variables, Lith. Math. J. 52 (2012), 214–221.
- [23] WANG, X. J.—WANG, S.—HU, S.—LING, J.—WEI, Y.: On complete convergence of weighted sums for arrays of rowwise extended negatively dependent random variables, Stochastics 85 (2013), 1060–1072.
- [24] WU, Q. Y.: Complete convergence for weighted sums of sequences of negatively dependent random variables, J. Probab. Stat. 2011 (2011), Article ID 202015.
- [25] XU, H.—TANG, L.: Some convergence properties for weighted sums of pairwise NQD sequences, J. Inequal. Appl. 2012 (2012), Article ID 255.
- [26] XU, H.—TANG, L.: Strong convergence properties for ψ-mixing random variables, J. Inequal. Appl. 2013 (2013), Article ID 360.
- [27] ZHOU, X. C.—TAN, C. C.—LIN, J. G.: On the strong laws for weighted sums of ρ\*-mixing random variables,
   J. Inequal. Appl. 2011 (2011), Article ID 157816.

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